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Long term economic impact associated with childhood bullying victimisation

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1

Long term economic impact associated with childhood bullying victimisation

2

3 Abstract

Being bullied is associated with mental health problems in childhood, with increasing evidence of 4 5 persisting negative impacts, and increased mental health service use, into adulthood. There are also б impacts of bullying victimisation on employment, income and being in poverty, but little is known 7 about the long-term economic impacts. We therefore aimed to estimate the most important 8 economic consequences at age 50 of being bullied in childhood: to our knowledge this is the first study that does so. Using 1958 British birth cohort data collected in 1965, 1969, 1991, 2003 and 9 2008 (study samples size 7,323-9,242), we find substantial and durable individual and societal 10 economic impacts four decades after the childhood bullying occurred. Both men and women who 11 12 were bullied in childhood were less likely to be in employment and had accumulated less wealth in 13 the form of home-ownership or savings than participants who were not bullied. Individual earnings from paid employment were lower for women who were bullied in childhood. Frequent bullying in 14 childhood was also associated with higher societal employment-related costs for men and higher 15 health service costs for women. Our findings underline the importance of preventing bullying in 16 childhood and, as the consequences are so long-lasting and pervasive, supporting people still 17 18 experiencing the negative consequences in the decades that follow.

19

20 Keywords

21 UK; childhood bullying; long-term; economic impact; mental health; health service use; employment;
 22 wealth

24 Introduction

25

A substantial body of evidence in the UK and internationally shows that bullying victimisation is associated with mental health problems in childhood (e.g. Arseneault *et al.*, 2018). There is also increasing evidence of persisting negative impacts into the adult years (e.g. Wolke and Lereya, 2015; Arseneault, 2017). Relatedly, it is not surprising that childhood bullying victimisation is associated with increased mental health service use in childhood, adolescence, and early and mid-adulthood up to age 50 (Evans-Lacko *et al.*, 2016; Sourander *et al.*, 2009; Sourander *et al.*, 2016).

childhood bullying victimisation was associated with more difficulties keeping a job when aged 24 to 34 26 (Wolke et al., 2013). We similarly found that men who were bullied in childhood were more likely 35 36 to be unemployed at age 50 than their peers (Takizawa et al., 2014), using British birth cohort data (the National Child Development Survey; NCDS). In other analyses of the NCDS, bullied children had 37 lower incomes than their peers at ages 23 and 33, although not at age 42 (Brown and Taylor, 2008). 38 This study suggested that being bullied at school lowered wages earned during adulthood directly, as 39 well as indirectly through educational attainment. Finally, in the US cohort, bullied children were 40 41 more likely to be living in poverty when aged 24 to 26 than their non-bullied peers (Wolke et al., 2013). 42

43

As these findings suggest, being bullied in childhood has adverse long-term consequences, but little is known about the long-term *economic* impacts, whether individual or societal. We aimed to estimate the most important economic consequences at age 50 of being bullied at ages seven and eleven, looking at a range of individual impacts (employment status, earnings and wealth) and societal impacts (costs of health service use for mental health problems and employment-related costs to society). These outcomes are linked not only to individual financial wellbeing and potential

50	costs to the state but may have other implications since wealth, both savings and home-ownership,
51	is of key importance to individuals and society in a range of respects, including mental and physical
52	health, social care and pensions (Crossley and O'Dea, 2010; ; Dunatchik, 2016; Marmot et al., 2010;
53	Pierse <i>et al.,</i> 2016; Pollack <i>et al.,</i> 2007).
54	
55	Data and Methods
56	
57	Sample
58	Participants are from the NCDS, the 1958 British birth cohort (Power and Elliott, 2006). Information
59	was collected on 98% of all births in one week in 1958 in England, Scotland and Wales. During
60	follow-up rounds of data collection at ages seven, eleven and sixteen years, the sample was
61	augmented by 920 immigrants to the UK born in the study week, resulting in a total of 18,558 cohort
62	members. Our sample for analysis comprises all cohort members for whom we have complete data
63	on bullying in childhood and a range of outcomes at age 50: mental health service use (n=9,242);
64	employment (n=8,581); earnings (n=7,323); housing tenure (n=9,222) and savings (n=7,559). Data
65	used in our analysis were collected in 1965, 1969, 1991, 2003 and 2008.
66	
67	Ethical approval was obtained by NCDS via NHS Research Ethics Committees (RECs) or, prior to their
68	set-up in 1997, through internal review. Ethical approval for the biomedical survey was given by the
69	South East Multi-Centre REC.
70	
71	Measures
72	Assessment of bullying
73	Exposure to bullying was assessed via parental interviews when participants were aged seven and

eleven years. At each age, parents were asked if their child was bullied by other children 'never',

'occasionally' or 'frequently'. We combined responses from both interviews to create a three-level
indicator of exposure to childhood bullying: 0 = never bullied (never at both seven and eleven years);
1 = occasionally bullied (occasionally at either seven or eleven years); 2 = frequently bullied
(frequently at either seven or eleven years, or occasionally at both ages). Where only one parental
interview was available, responses from that interview were used, providing bullying assessments
for 86% of cohort members.

- 81
- 82

83 Individual economic impacts at age 50

We looked at three possible individual economic impacts of childhood bullying: economic status, 84 weekly earnings and wealth. Earnings penalties of bullying victimisation were based on comparisons 85 of self-reported weekly net individual earnings from paid employment at age 50. In NCDS, this 86 87 question was only asked of participants in paid employment and so did not include people in selfemployment, unemployed or economically inactive. Measures of wealth at age 50 were housing 88 tenure and savings. Housing tenure was categorised as either owner-occupation or renting, the 89 latter comprising both private and social renting. For savings, we created a categorical variable of 90 'no-to-low' savings, defined as savings between zero and £392.92, 'low-to-median' savings and 91 above median savings. The upper level for no-to-low savings (£392.92) was based on the PSE2012 92 (Poverty and Social Exclusion in the UK) Individual Deprivation Measure (Gordon, 2017) deflated to 93 94 2008 values (i.e. the year when NCDS cohort members were age 50), using the Retail Price Index.

95

96 Societal economic impacts at age 50

We considered two measures of *societal* economic impact: costs associated with health service use
for mental health problems and employment-related costs to society. In estimating these, we used
2008 cost levels throughout, to match with the date of the age 50 data collection.

101 Health service use for mental health conditions

As the majority of mental health care occurs outside of specialty 'mental health' care settings 102 (Brown et al., 2014; Evans-Lacko, 2017), and there is a widespread reluctance among individuals to 103 report seeking help for mental health problems (Clement et al., 2015), we included service use for 104 mental health conditions occurring in both specialty mental health and general health settings. We 105 106 looked at three types of health service use reported by participants at age 50: specialty mental health service use over the previous four years, general hospital outpatient service use and general 107 108 hospital inpatient service use over the previous eight years. To estimate costs associated with hospital inpatient service use, we multiplied the number of hospital inpatient days reported by the 109 participant over the previous eight years by the World Health Organization mean cost estimate for 110 the UK per hospital inpatient bed-day in 2008 (WHO, 2011). For frequency of outpatient service use 111 in the previous eight years, possible responses included: none, one or two, three to five, six to ten 112 113 and more than ten. The median number of visits for each respondent was multiplied by the mean unit cost per outpatient visit in 2008 (Curtis, 2008). As frequency of specialty mental health service 114 use at age 50 is unavailable in NCDS, we estimated annual service use frequency and associated 115 costs according to national averages stratified by gender from the 2000 Adult Psychiatric Morbidity 116 Survey for the sub-sample aged 42 to 50. To enable aggregation of costs across all three types of 117 118 health service use, annual specialty mental health costs were multiplied by eight to give estimated 119 eight-year costs.

120

121 Employment-related societal costs at age 50

We estimated employment-related costs to society using the human capital approach, which has been widely used in other economic evaluations (e.g. Park *et al.*, 2014). Costs were calculated for individuals in employment, self-employment, unemployment or economically inactive through temporary or permanent sickness or disability. We considered full-time employment to be 35 or more hours per week and applied the national minimum hourly wage in 2008 (£5.73 an hour) to

hours worked per week less than the 35-hour full-time equivalent. This was then multiplied by 48 on 127 the assumption of 48 working weeks in one year. Finally, we estimated aggregate annual societal 128 costs of being bullied in childhood by multiplying per-person societal costs by the estimated 129 numbers aged 50 in 2008 who had been bullied in childhood. This was estimated using prevalence of 130 being frequently bullied in childhood from our NCDS sample and mid-year estimates of the UK 131 132 population aged 50 in 2008 (ONS, 2012). 133 134 Covariates Childhood confounders 135 We controlled for childhood confounders known to be associated with bullying, and with the 136 outcomes under study (Takizawa et al., 2014; Evans-Lacko et al., 2016). Childhood IQ was assessed 137 at age eleven using a standardized 80-item general ability test (Douglas, 1964). Scales of childhood 138 139 emotional and behavioural problems were derived from teacher ratings on the Bristol Social Adjustment Guides (Stott, 1969) at ages seven and eleven years. We used the mean of scores across 140 ages seven and eleven years where both measures were available, and single-age measures for the 141 remainder of the sample. Family social class in childhood was classified on the basis of the father's 142 occupation when the sample member was aged seven years, and categorised as 'I and II' 143 144 professional/managerial/technical, 'IIINM' other non-manual, 'IIIM' skilled manual, and 'IV and V' 145 unskilled manual (OPCS, 1980). Childhood adversity was assessed from both prospective and 146 retrospective reports. Prospectively, information collected from parents and teachers was used to 147 create an eight-item scale of low parental involvement and activity with the child at ages seven and eleven (Power and Elliott, 2006). Retrospectively at age 45, participants completed a 16-item 148 149 questionnaire about their exposure to a range of childhood adversities including poverty, parental 150 mental ill-health and drug/alcohol problems, family conflict, and physical and sexual abuse (Rosenman, 2004), but not childhood bullying. 151

152

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154	
155	Adult covariates
156	In follow-on analyses, we also controlled for covariates in adulthood (age 33) known to be
157	associated with both being bullied in childhood (Brown and Taylor, 2008; Takizawa et al., 2014) and
158	with the economic impacts. These covariates were partnership status, highest educational
159	qualification and psychological distress as measured by the Malaise Inventory, a nine-item
160	questionnaire with validity in population samples (Rodgers, 1999). Data on these covariates were
161	collected after the bullying victimisation took place and before the economic impacts at age 50, so
162	could indicate ways in which childhood bullying victimisation might affect economic outcomes at age
163	50.
164	
165	Analysis
166	We report descriptive statistics to characterise the sample by bullying victimisation and by gender,
167	with Chi-squared tests and tests of means used to determine any group differences. In comparing
168	economic outcomes by bullying victimisation, we considered two sets of regression models. Firstly,
169	bivariate analyses compared economic outcomes between participants who were occasionally or
170	frequently bullied in childhood with those who were not. The second set of regression models
171	
	additionally controlled for childhood confounders. For linear outcome variables – earnings, health
172	additionally controlled for childhood confounders. For linear outcome variables – earnings, health service costs and employment-related costs - we utilised two-part generalised linear models (GLM)
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172 173 174	additionally controlled for childhood confounders. For linear outcome variables – earnings, health service costs and employment-related costs - we utilised two-part generalised linear models (GLM) (Mullahy, 1998). As the dependent variables may have skewed distributions, we used a modified Park test (Manning and Mullahy, 2001), to select the most appropriate distribution for the purposes
172 173 174 175	additionally controlled for childhood confounders. For linear outcome variables – earnings, health service costs and employment-related costs - we utilised two-part generalised linear models (GLM) (Mullahy, 1998). As the dependent variables may have skewed distributions, we used a modified Park test (Manning and Mullahy, 2001), to select the most appropriate distribution for the purposes of estimation. The marginal effect of bullying status on each outcome was estimated for each
172 173 174 175 176	additionally controlled for childhood confounders. For linear outcome variables – earnings, health service costs and employment-related costs - we utilised two-part generalised linear models (GLM) (Mullahy, 1998). As the dependent variables may have skewed distributions, we used a modified Park test (Manning and Mullahy, 2001), to select the most appropriate distribution for the purposes of estimation. The marginal effect of bullying status on each outcome was estimated for each regression model, generating estimates of mean cost differences for people who were occasionally
172 173 174 175 176 177	additionally controlled for childhood confounders. For linear outcome variables – earnings, health service costs and employment-related costs - we utilised two-part generalised linear models (GLM) (Mullahy, 1998). As the dependent variables may have skewed distributions, we used a modified Park test (Manning and Mullahy, 2001), to select the most appropriate distribution for the purposes of estimation. The marginal effect of bullying status on each outcome was estimated for each regression model, generating estimates of mean cost differences for people who were occasionally bullied and frequently bullied compared to people who were never bullied. For categorical outcome

179	models. As health service use, employment rates, earnings and wealth tend to differ by gender
180	(Evans-Lacko et al., 2016; ONS, 2013; ONS, 2016), we estimated economic impacts separately for
181	women and men.
182	
183	Finally, we explored potential mediators of these economic impacts, by additionally controlling for
184	adult (age 33) covariates - partnership status, highest educational qualification and psychological
185	distress - in our regression models. Covariates that reduce the size of the effect can be considered as
186	potentially plausible explanations for the observed associations.
187	
188	All tests of statistical significance used robust standard errors. A significance level of 0.05 was used
189	as the criterion to determine statistical significance and 0.10 to indicate marginal significance. We
190	conducted the analyses using Stata 14.2 (StataCorp, 2015).
191	

192 Attrition

193 To investigate whether attrition was associated with the variables in our analysis, we explored the savings variable as this had the highest proportion of missing data. Descriptive data showed that 194 195 data non-availability because of attrition and/or missingness was unrelated to childhood bullying but was associated with the childhood confounders, except adversity (supplementary table available 196 from authors on request). To account for potential attrition bias, the analyses incorporated inverse 197 probability weights to address this differential sample attrition. These were derived from logistic 198 regression analyses predicting availability of complete data on childhood bullying and each outcome 199 200 at age 50, including all the childhood confounder variables except adversity.

201

202

204	
205	Results
206	Childhood and early adult correlates of bullying victimisation
207	Rates of childhood bullying victimisation were higher among men than women (table 1). Being
208	occasionally or frequently bullied in childhood was associated with a range of potential childhood
209	confounders. Both women and men who were bullied in childhood had lower general ability test
210	scores, and elevated levels of internalising and externalising problems in childhood compared to
211	those who were never bullied. They were more likely to come from skilled or unskilled manual social
212	class backgrounds than those who were never bullied. They were at higher risk of experiencing other
213	forms of childhood adversity. Table 1 also shows that being occasionally or frequently bullied in
214	childhood was associated with having lower educational qualifications and higher psychological
215	distress at age 33 for both men and women and less likelihood of having a partner at age 33 for men.
216	
217	<table 1="" about="" here=""></table>
218	
219	Economic impacts at age 50: women
220	Our bivariate analysis showed that, for women, being occasionally or frequently bullied in childhood
221	was significantly associated with worse economic outcomes at age 50 in every domain (table 2). The
222	addition of childhood confounders reduced the odds ratios (OR), relative risk ratios (RRR), or costs of
223	poorer economic outcomes associated with bullying alone, in some cases substantially. However,
224	even after accounting for these childhood factors there were still significant economic consequences
225	at age 50 of being frequently bullied in childhood. The main exception was societal employment
226	costs where differences in costs by bullying for women were accounted for by differences in the
227	childhood confounders.
228	<table 2="" about="" here=""></table>
229	

230	Controlling for childhood confounders, women who were occasionally or frequently bullied in
231	childhood had marginally higher odds of 1.34 (95% CI=0.97, 1.84) and 1.39 (95% CI=0.94, 2.06)
232	respectively of being unemployed or economically inactive due to sickness or disability compared to
233	women who were never bullied. Women who were frequently bullied in childhood had lower net
234	weekly earnings from paid employment (on average £22.74 a week lower at 2008 values) and less
235	likelihood of owning a property at age 50 (OR=0.76; 95% CI=0.57, 1.01) than women who were not
236	bullied, even controlling for childhood confounders. Additionally, women who were frequently
237	bullied in childhood had higher odds of having accumulated only no-to-low (RRR=1.68; 95% CI=1.23,
238	2.29) or low-to-median (RRR=1.80; 95% CI=1.37, 2.38) savings by mid-life compared to those who
239	were never bullied.
240	
241	Focusing on societal impact, being frequently bullied in childhood was associated with £717 higher
242	health service costs for mental health conditions over an eight-year period at mid-life among
243	women, even controlling for childhood confounders. Aggregated to the estimated affected
244	population, this was an estimated £4.5 million annually.
245	
246	Economic impacts at age 50: men
247	Bivariate analysis showed that for men, being frequently bullied in childhood was associated with
248	worse economic outcome at age 50 in every domain and occasionally bullied with worse outcomes
249	in most domains (table 3). Similarly to women, the addition of childhood confounders reduced the
250	odds or costs of the economic outcomes associated with being bullied, although significant
251	economic impacts were still observed in many areas. The exceptions were weekly income and health
252	service costs where childhood confounders explained the differences seen.
253	
254	Controlling for childhood confounders, men who were frequently bullied in childhood were at

increased risk (OR=1.49; 95% CI =1.04, 2.13) of being unemployed or economically inactive due to

256	sickness or disability compared to men who were never bullied. We did not observe differences in
257	weekly net earnings at mid-life for men who experienced bullying in childhood. Men who were
258	frequently bullied in childhood had lower odds of owning a property at age 50 compared to their
259	non-bullied peers (OR=0.74; 95% CI=0.56, 0.97), and marginally higher risk of having no-to-low
260	savings (RRR=1.31; 95% CI=0.97, 1.78).
261	
262	<table 3="" about="" here=""></table>
263	
264	Moving on to societal impacts, we found employment-related societal economic impacts were
265	higher for men who were frequently bullied in childhood compared to those who were never
266	bullied, with extra costs to society of £271 annually, even controlling for childhood confounders.
267	Applied to the estimated affected population in 2008, this represents an aggregate cost to society of
268	an estimated £17.9 million a year.
269	
270	
271	Adult mediators of economic impacts
272	We then turned to possible mediators of the economic impacts of being bullied in childhood:
273	partnership status, highest educational qualification and psychological distress, all at age 33 (Table
274	4).
275	
276	< TABLE 4 ABOUT HERE >
277	
278	Adding these adult covariates to the model reduced the earnings difference for women frequently
279	bullied in childhood compared to never bullied by about a quarter and reduced the odds ratio of
280	being unemployed or economically inactive for women who were occasionally bullied compared to
281	never bullied. The adult covariates did not appear to have much effect on any of the other individual

outcomes for women, although they did reduce estimates of the societal costs of health service use 282 283 for mental health conditions by just under a third. For men, the addition of adult covariates appeared to make little difference to odds of being unemployed/economically inactive due to long-284 term sickness or disability, odds of owner-occupation nor odds of having no-to-low savings when 285 aged 50, although societal employment-related costs were reduced by a third. These results suggest 286 287 that partnership status, highest educational qualification and psychological distress at age 33 are potential explanations for some of the economic impacts at age 50 of being bullied in childhood. 288 289 However, in most cases, economic impacts were still observed even after additionally controlling for these factors. 290

291

292 Discussion

To our knowledge, this is the first study that looks at the long-term economic consequences of 293 294 childhood bullying victimisation. We find substantial and durable individual and societal economic impacts at mid-life of being bullied in childhood. Four decades after the bullying occurred, both men 295 and women who were bullied in childhood were less likely to be in employment and had 296 accumulated less wealth in the form of home-ownership or savings than participants who were not 297 bullied. Additionally, women who were bullied in childhood had lower earnings from paid 298 299 employment. We also identified societal economic impacts associated with childhood bullying: 300 frequent bullying in childhood was associated with higher employment-related costs at age 50 for 301 men and higher health service costs at age 50 for women. The latter is consistent with the higher 302 mental health service use seen by adults who were bullied as children in other research (Evans-Lacko et al., 2016; Sourander et al., 2016). 303

304

The reasons for lower earnings and higher unemployment among those who were bullied during childhood are likely to be similar. Both are associated with lower educational attainment, and bullied children have lower educational attainment than their non-bullied peers (Brown and Taylor, 2008)).

Bullied children miss more school than non-bullied children (Brown et al., 2011), which may be 308 309 because being bullied at school results in reluctance to attend and/or because of the contemporaneous mental and physical ill-health associated with bullying. Bullied children may find it 310 harder to concentrate when they are in school, again either directly because of the bullying or 311 because of the associated health problems, psychological difficulties in particular. Furthermore, 312 313 bullying victimisation may lead to cognitive problems which can affect educational attainment (Takizawa et al., 2014). Low self-esteem associated with bullying victimisation (Smokowski and 314 315 Holland, 2005) is another potential mechanism by which educational attainment may be affected (Waddell, 2006). 316

317

Many of the adulthood consequences of childhood bullying victimisation may also directly impact on 318 employment and earnings. The lower self-esteem and confidence of adults who were bullied as 319 320 children (Waddell, 2006) plausibly impacts on job-seeking or promotion-seeking, working hours, or performance at work. Another potential pathway is that being bullied may alter physiological 321 responses to stress (Ouellet-Morin et al., 2011) which may lead to withdrawal from the labour 322 market or reduction in working hours. Poorer mental health is associated with higher unemployment 323 and lower income, as are physical health problems. For example, the poorer mental and physical 324 325 health reported by adults who have been bullied as children often stands in the way of work (Allison 326 et al., 2009) and in our analysis the costs associated with non-employment include those who are 327 permanently or temporarily unable to work because of sickness. Relationship status is associated 328 with employment and people who are bullied in childhood are less likely to be in a relationship than their non-bullied peers (Takizawa et al., 2014). In our analysis exploring the role of educational 329 330 attainment, mental health and relationship status further, we found that these factors at age 33 appear to be potential explanations in the earnings differences for frequently bullied women and 331 the employment costs for frequently bullied men. 332

333

We found wealth penalties of being bullied in two domains: housing and savings. Wealth can be 334 considered a measure of accumulation of assets over the life-course. Our findings thus show not 335 only the economic consequences of bullying four decades after being bullied in childhood, but also 336 point to accumulative, life-course, economic consequences. There are significant implications of 337 having low wealth in mid-life. Wealth can provide protection against both present and future 338 339 financial shocks associated with, for example, unemployment, serious illness or relationship breakdown, and hence provide a measure of financial resilience. The cut-off point for low savings we 340 341 use is drawn from a robust scale of deprivation items (Gordon, 2017). Savings above this level can provide some financial protection against unexpected, but necessary, expenses. Without this, there 342 is higher risk of debt, which could become unmanageable, or going without necessities. 343 344 Savings wealth is linked to other types of wealth which provide financial protection: for example, 345 346 those with higher savings are more likely to own their home and to contribute to pension schemes (Crossley and O'Dea, 2010). As well as having individual economic impacts, savings wealth has other 347 implications, both individual and societal. One currently quite high-profile example is paying for 348 social care. For example, at present in England, savings wealth determines financial eligibility for 349 state funding for social care. Furthermore, research in England has shown a relationship between 350 351 level of savings and unmet need for services (Dunatchik, 2016). Wealth, especially housing wealth, is an important factor for better physical and mental health (Pierse *et al.*, 2016; Pollack *et al.*, 2007) 352 and can be considered one of the most important contributors to socio-economic health inequalities 353 (Marmot et al., 2010; Pollack et al., 2007;). 354

355

Accumulation of wealth results from being consistently in adequately paid employment over the longer term, and thus some of the potential ways in which bullying may impact on earnings and/or employment are also relevant to wealth. Another factor may be inheritance, which plays a role in both home-ownership and wealth. However, our analyses controlled for family socio-economic class,

one measure of family wealth and hence inheritance, and the association of home-ownership with 360 bullying still remains. People who have been bullied in childhood are less likely to be in relationships 361 as adults (Takizawa et al., 2014), and partnership status is strongly associated with home-ownership 362 (Thomas and Mulder, 2016). However, once again, the negative association between home-363 ownership and being bullied remains even when controlling for partnership status age 33. 364 365 The possible ways in which bullying in childhood may impact on mental ill-health in adulthood even 366 367 four decades after its occurrence have been reviewed elsewhere (Evans-Lacko et al., 2016; Sourander et al., 2016; Takizawa et al., 2014; Wolke and Lereya, 2015) and include long-term trauma 368 of early adverse experiences increasing vulnerabilities to later mental health problems (Shonkoff et 369 al. 2009), interpersonal processes (Kendler et al., 2003), and physiological responses (Ouellet-Morin 370 et al., 2011). There may be other common factors for being exposed to bullying and mental health 371 372 problems which we did not control for.

373

Mental health problems associated with bullying might translate into higher health service use 374 among bullied children and there is certainly some evidence for this (Evans-Lacko et al., 2016; 375 Sourander et al., 2016). However, the relationship between health service use for mental health 376 377 problems and the existence of those problems is by no means straightforward. Under- or mis-378 diagnosis, stigma, sub-threshold mental health problems and lack of available and/or suitable 379 services are some of the reasons why the presence of a mental health problem does not necessarily result in accessing services for that problem; it is telling that differences in health service costs, 380 although significantly reduced, remain for women even after controlling for one measure of 381 382 psychological distress in adulthood. Nonetheless our previous research found a significant 383 relationship between bullying victimisation in childhood and mental health service use in adulthood, 384 although at age 50 this was only the case for women (Evans-Lacko et al., 2016). Here we have extended that analysis, factoring in intensity of service use by considering both frequency and type 385

of health service use, and then estimating different costs by gender. Using this method, we found
that health service costs for women who were frequently bullied were more than £700 higher than
for their non-bullied peers, with no significant differences in costs for men. Reasons for this gender
difference may be to do with the greater help-seeking among women with mental health problems
seen in other research (Evans-Lacko *et al.*, 2014).

391

There were other gender differences observed in our study. Women who were bullied in childhood 392 393 had lower earnings than their non-bullied peers, whereas bullied men did not. This is likely to be due to differences in hours worked per week, which is a more important determinant of earnings for 394 women than for men (Blundell et al., 2013). Both men and women had higher odds of having no-to-395 low savings if they had been frequently bullied in childhood. However, women who were frequently 396 bullied in childhood also had higher odds of having low-to-median savings whereas there were no 397 differences for men. Both men and women were less likely to be in employment if they had been 398 bullied in childhood, but it was men for whom there were associated societal costs. Overall, we 399 found that childhood bullying appeared to have a differential, and worse, economic impact in later 400 life for women than for men. This may be because women and men are differently affected by 401 childhood bullying in the first place, and/or because the negative impacts of being bullied manifest 402 403 differently for women and men because of gender differences in the wider socio-economic 404environment. Although there are no other studies on the long-term economic consequences of 405 being bullied in childhood, research on childhood maltreatment concluded that women were more 406 vulnerable than men to the long-term economic impacts of childhood abuse or neglect (Currie and Widom, 2010). 407

408

There are a number of limitations to our study. Attrition by age 50 and missing data reduces the
number of cases for some of our outcomes. However, we controlled for effects of selective attrition
by utilising inverse probability weights in the analyses. Information on exposure to bullying was

assessed via parental interviews in NCDS so may underestimate incidence, as not all childhood
bullying is known about by parents. Shakoor and colleagues (2011) found 'modest' (52-56%)
agreement between mother and child reports of bullying at this age but similar associations
between child emotional and behavioural problems and bullying victimisation, regardless of
informant. Whilst data from more than one informant is the preferred option, in the absence of child
self-reports, mothers' reports can be considered a viable alternative.

418

419 Another potential limitation of our study is that there may be recall bias, in particular for service use, although self-report is considered an acceptable method for collecting service use data (Patel et al., 420 2005). Self-reports of earnings can also be subject to recall or estimation bias, although the NCDS 421 question asked about very recent, and short, timeframes. Reported earnings can also be sensitive to 422 social desirability bias, but the self-completion or computer-assisted methods used in NCDS reduces 423 424 this risk considerably. Earnings are limited to earnings from paid employment, an approach that has been used in other similar research using NCDS. Use of average frequencies of service use 425 interpolated over the time frame and average national costs for estimation of health service costs 426 may have resulted in under- or over- estimates of actual costs. Using the minimum wage to estimate 427 societal employment-related costs is also likely to be an underestimate. Furthermore, although we 428 429 have looked at both individual and societal impacts, and in a number of domains, there may be 430 other economic consequences of bullying which we have not estimated here.

431

The strengths of our analyses are that they are based on a large nationally representative cohort dataset with rich data on childhood bullying victimisation and on a range of outcomes collected through to age 50 to which we can apply individual and societal costs. We have also been able to control for key confounders, showing that our analyses are robust. Using these data, our study is the first to estimate individual and societal costs in mid-life in a number of spheres of bullying in childhood.

438	
439	
440	Conclusions
441	Bullying in childhood is widespread, with consistent evidence of negative impacts across diverse
442	samples in terms of age of being bullied, time, geographic location and culture. Internationally, one
443	in three children report having been bullied at some point in their lives, with 10% to 14%
444	experiencing chronic bullying lasting for more than six months (World Health Organization, 2012).
445	Bullying was reported by 34% to 46% of school children in England in recent surveys (Department of
446	Health, 2015), similar to the prevalence in our NCDS sample. Vulnerable children are at higher risk
447	of being bullied (Arseneault et al., 2018; Chatzitheochari et al., 2014; Gower et al., 2015; McMahon,
448	2010; Woods et al., 2009). That it has personal and economic consequences for both individuals and
449	society four decades later should raise significant concerns.
450	
451	By showing the economic consequences for individuals and society of being bullied in childhood, our
452	study further underlines the importance of investing in the implementation of effective policies and
453	practices to prevent or reduce bullying in schools and thus mitigate negative outcomes and costs. Of
454	course, tackling bullying is particularly a concern because of the quality of life effects on those who
455	are bullied. There is evidence of effective anti-bullying initiatives (see for example the meta-analysis
456	by Ttofi and Farrington, 2011) and indeed cost-effective initiatives, many of which are low cost (from
457	£8-16 per person per year) (Beecham <i>et al.,</i> 2011; McDaid <i>et al.,</i> 2017). This compares to our
458	estimated annual societal costs per person of being bullied of £90 for women and £271 for men at
459	age 50 in 2008. As risk of being bullied is higher among already disadvantaged children and the
460	poorer economic outcomes seen here compound that disadvantage in later life, anti-bullying
461	initiatives also have a potentially important role to play in supporting vulnerable, disadvantaged
462	children and reducing inequities in later life.

464 As the consequences of childhood bullying are long-lasting and pervasive, it is imperative that 465 response in this area should be multi-faceted. It needs to prevent bullying, to provide support for children (and adults) who are currently being bullied, and to provide support for people who were 466 bullied previously (whether in childhood or subsequently) and who are still experiencing the 467 negative consequences of it in the years that follow. Better integration of mental health and anti-468 bullying provision could prevent young victims developing long-term mental health problems, and 469 the costs associated with this. Given the potential mediating role in adulthood of psychological 470 distress, and educational qualifications and relationship status, supporting adults bullied as children 471 might usefully include relationship and other social support, access to adult education and support 472 with mental health problems. 473

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Table 1. Descriptive statistics of sample childhood socio-demographic characteristics by bullying and gender

		Female (n=4,702)			Male (n=4,540)	
Bullied in childhood	Never % (N) 61.6 (2,895)	Occasionally % (N) 26.2 (1,231)	Frequently % (N) 12.3 (576)	Never % (N) 55.7* ¹ (2,527)	Occasionally % (N) 28.5 ^{*1} (1,296)	Frequently % (N) 15.8 ^{*1} (717)
CHILDHOOD CONFOUNDERS						
General ability test score (mean; SD)	48.0 (14.9)	44.7* (15.3)	42.2* (15.2)	45.5 (15.3)	43.3* (16.0)	41.8* (16.2)
Internalising score (mean; SD)	1.75 (0.85)	1.90* (0.88)	2.01* (0.97)	1.89 (0.86)	2.06* (0.94)	2.22* (0.98)
Externalising score (mean; SD)	1.76 (0.84)	1.84* (0.90)	1.94* (0.93)	1.97 (0.97)	2.08* (1.03)	2.22* (1.08)
Parental social class % (N)		Č				
Skilled professional/ managerial/technical	23.6 (676)	18.7*(230)	16.2*(93)	25.3 (634)	19.4*(250)	15.6*(112)
Skilled non-manual	10.5 (302)	10.0*(123)	7.64*(44)	9.9 (248)	11.4*(147)	8.7*(62)
Skilled manual	42.7 (1,224)	43.7*(536)	44.3*(255)	40.7 (1,020)	42.1*(543)	50.5*(362)
Unskilled manual	23.3(668)	27.6*(339)	31.9*(184)	24.1 (604)	27.1*(349)	25.2*(181)
Childhood adversity score (mean; SD)	1.51 (2.29)	1.74* (2.38)	1.90* (2.38)	1.19 (1.80)	1.41* (2.09)	1.52* (2.08)
Low parental involvement score (mean; SD)	0.90 (1.25)	1.11* (1.40)	1.15* (1.40)	0.93 (1.30)	1.11* (1.40)	1.42* (1.67)
ADULT COVARIATES (AGE 33)		\mathcal{O}				
Partner % (N)	83.5 (2,098)	80.2 (865)	83.6 (418)	83.1 (1,754)	80.5*(860)	75.1*(432)
Highest educational qualification % (N)	R					
None	9.2(232)	12.5*(134)	17.6*(88)	7.4 (156)	8.7*(93)	13.9*(80)
CSE, O Level or equivalent	49.6(1,256)	52.9*(569)	54.5*(273)	39.4 (835)	42.8*(458)	41.2*(237)
A Level or equivalent	27.3(691)	23.6*(254)	20.6*(103)	36.0 (763)	33.2*(355)	33.9*(195)
Degree or higher	13.9(351)	11.1*(119)	7.4*(37)	17.3 (368)	15.4*(165)	11.1*(64)
Psychological distress (mean; SD)	1.07 (1.53)	1.36* (1.73)	1.51* (1.86)	0.58 (1.15)	0.74* (1.29)	0.90* (1.46)

*p < .05; reference group is never bullied except (1) where reference group is women

Table 2 Association between being	$_{2}$ bullied in childbood and individual and societal economic impact for women at age 50	
Table 2. Association between being	, builled in childhood and individual and societal economic impact for women at age so	

INDIVIDUAL ECONOMIC IMPACTS	Bullied at age 7 & 11		Economic impact: bivariate results	Economic impact controlling for childhood confounders ¹		
Economic status		%	Odds ratio (95% CI)			
Unemployed or economically inactive	Never (n=2,544)	6.8	-	-		
	Occasionally (n=1,099)	8.7	1.44* (1.09, 1.91)	1.34~ (0.97, 1.84)		
	Frequently (n=504)	11.9	1.85* 1.32, 2.58)	1.39~ (0.94, 2.06)		
Earnings		Mean	Earnings diffe	rence (95% Cl)		
Mean weekly earnings (£) from paid employment	Never (n=2,015)	301.25	-	-		
	Occasionally (n=883)	281.27	-21.05* (-37.76, -4.34)	-8.72 (-25.92, 8.48)		
	Frequently (n=386)	258.50	-41.92* (-60.89, -22.96)	-22.74* (-42.05, -3.43)		
Housing tenure		%	Odds ratio	o (95% CI)		
Not owner occupier	Never (n=2,889)	14.3	-			
	Occasionally (n=1,230)	18.3	0.76* (0.63, 0.93)	0.87 (0.69, 1.10)		
	Frequently (n=573)	20.6	0.65* (0.51, 0.83)	0.76~ (0.57, 1.01)		
Savings		%	Relative risk ratio (95% CI)			
No-to-low savings compared to above median	Never (n=2,281)	20.2	-	-		
savings	Occasionally (n=975)	25.6	1.51* (1.23, 1.85)	1.25~ (0.99, 1.58)		
	Frequently (n=432)	30.1	2.03* (1.55, 2.67)	1.68 * (1.23, 2.29)		
Low-to-median savings compared to above	Never (n=2,281)	29.6	-	-		
median savings	Occasionally (n=975)	30.2	1.27* (1.05, 1.53)	1.67 (0.95, 1.44)		
	Frequently (n=432)	33.3	1.71* (1.33, 2.21)	1.80* (1.37, 2.38)		
SOCIETAL ECONOMIC IMPACTS		Mean	Cost differe	nce (95% CI)		
Eight-year health service costs for mental health problems (£)	Never (n=2,884)	2516.76	-			
Č	Occasionally (n=1,222)	2814.83	422.96~ (-16.28, 862.20)	371.06 (-63.65,805.77)		
	Frequently (n=570)	3470.11	1,040.33 * (546.94, 1533.73)	717.08* (244.39, 1,189.76)		
Mean annual societal employment cost (£)						
	Never (n=2,527)	2175.18	-			
	Occasionally (n=1,093)	2284.62	243.27~ (-1.58, 488.12)	68.73 (-173.01, 310.47)		
	Frequently (n=499)	2549.22	463.16* (116.88, 809.44)	123.45 (205.33, 452.23)		

*p < .05; ~=p <0.10; Significant and marginal associations are highlighted in bold. 1. Controlling for family social class, adversity, low parental involvement, childhood IQ, childhood emotional and behavioural problems

INDIVIDUAL ECONOMIC IMPACTS	Bullied at age 7 & 11		Economic impact: bivariate results	Economic impact controlling for childhood confounders ¹		
Economic status	%		Odds ratio (95% CI)			
Unemployed or economically inactive	Never (n=2,472)	7.7	-	-		
	Occasionally (n=1,264)	9.2	1.21 (0.92, 1.58)	0.90 (0.64, 1.27)		
	Frequently (n=698)	14.0	2.06* (1.56, 2.72)	1.49 * (1.04, 2.13)		
Earnings		Mean	Earnings diffe	erence (95% CI)		
Mean weekly earnings (£) from paid employment	Never (n=1,622)	539.72		-		
	Occasionally (n=849)	524.38	-23.49(-68.27, 21.28)	-7.72 (-49.10, 34.66)		
	Frequently (n=448)	464.21	-70.34* (-105.90, -34.80)	-8.55 (-46.98, 29.88)		
Housing tenure		%	Odds rat	io (95% CI)		
Not owner occupier	Never (n=2,528)	15.0	-	-		
	Occasionally (n=1,292)	17.0	0.82* (0.67, 1.00)	0.99 (0.78, 1.26)		
	Frequently (n=710)	22.5	0.57* (0.46, 0.72)	0.74 * (0.56, 0.97)		
Savings	%		Relative risk ratio (95% CI)			
No-to-low savings compared to above median savings	Never (n=2,130)	16.8		-		
	Occasionally (n=1,120)	20.3	1.35* (1.09, 1.67)	1.12 (0.87, 1.45)		
	Frequently (n=621)	26.3	1.90* (1.49, 2.43)	1.31~ (0.97, 1.78)		
Low-to-median savings compared to above median	Never (n=2,130)	36.0	-	-		
savings	Occasionally (n=1,120)	35.7	1.12 (0.94, 1.33)	1.00 (0.82, 1.21)		
	Frequently (n=621)	35.1	1.22~ (0.98, 1.51)	1.11 (0.87, 1.40)		
SOCIETAL ECONOMIC IMPACTS		Mean	Cost differe	ence (95% CI)		
Eight-year health service costs for mental health	Never (n=2,515)	2270.53	-	-		
problems (£)	Occasionally (n=1,288)	2777.60	554.80* (65.77, 1,043.84)	-97.2 (-510.31, 315.92)		
	Frequently (n=714)	3199.91	1,141.25* (348.19, 1,934.31)	481.03 (-116.11, 1078.17)		
Mean annual societal employment cost (£)						
	Never (n=2,459)	921.98		-		
	Occasionally (n=1,253)	1117.06	189.19~ (-35.31, 407.69)	14.64 (-183.38, 212.66)		
	Frequently (n=693)	1588.72	728.31* (417.70, 1038.93)	270.70 * (5.66, 535.74)		

Table 3. Association between being bullied in childhood and individual and societal economic impact for men at mid-life

*p < .05; ~=p <0.10; Significant and marginal associations are highlighted in bold. 1. Controlling for family social class, adversity, low parental involvement, childhood IQ, childhood emotional and behavioural problems

Table 4: Association between being bullied in childhood and individual and societal economic impact for women and men at mid-life additionally controlling for adult covariates

		Women	Men	
INDIVIDUAL IMPACTS	Bullied at age	Childhood confounders and	Childhood confounders and	
	7 & 11	Adult covariates	Adult covariates	
Economic status		Odds ratio (95% CI)		
Unemployed or economically inactive	Occasionally	1.07	_	
		(0.73, 1.56)		
	Frequently	1.22	1.37	
		(0.78, 1.91)	(0.92, 2.06)	
Earnings		Cost difference (95% CI)		
Mean weekly earnings (£) from paid employment	Occasionally	-	· · · · · ·	
	Frequently	-16.58~ (-55.22, 2.12)		
Housing tenure		Odds ratio	Odds ratios (95% CI)	
Not owner occupier	Occasionally	-	-	
	Frequently	0.84	0.78	
		(0.61, 1.56)	(0.57, 1.07)	
Savings		Relative risk	ratio (95% CI)	
No-to-low savings compared to above median savings	Occasionally	1.16	-	
		(0.90, 1.50)		
	Frequently	1.55*	1.22	
		(1.08, 2.22)	(0.87, 1.73)	
Low-to-median savings compared	Occasionally		-	
to above median savings	Frequently	1.75*		
		(1.30, 2.36)	-	
SOCIETAL ECONOMIC IMPACTS				
		Cost differe	Cost difference (95% CI)	
Eight-year health service costs for mental health problems (£)	Occasionally	<u> </u>	-	
	Frequently	520.20*		
		(39.79, 1,000.61)	-	
		Cost difference (95% CI)		
Mean annual societal employment cost (£)	Occasionally	-	-	
	Frequently		£173.92	
		-	(-93.03,440.87)	

*=p< 0.05; ~=p <0.10; Significant and marginal associations are highlighted in bold.

Controlling for family social class, adversity, neglect, childhood IQ, childhood emotional and behavioural problems and partnership status, highest educational qualification and psychological distress age 33

Long term economic impact associated with childhood bullying victimisation

Highlights

- First study to explore long-term economic effects of being bullied in childhood
- Sizeable individual and societal economic impacts 4 decades after bullying occurred
- Adults bullied in childhood less likely to be in employment in mid-life
- Childhood bullying impacted on individual's mid-life income and accumulated wealth
- Higher societal employment costs for men and higher health service costs for women