



Eric Neumayer

## National carbon dioxide emissions : geography matters

Originally published in Area, 36 (1). pp. 33-40 © 2004 Blackwell Publishing.

You may cite this version as:

Neumayer, Eric (2004). National carbon dioxide emissions : geography matters [online]. London: LSE Research Online.

Available at: <http://eprints.lse.ac.uk/archive/00000602>

Available online: February 2006

LSE has developed LSE Research Online so that users may access research output of the School. Copyright © and Moral Rights for the papers on this site are retained by the individual authors and/or other copyright owners. Users may download and/or print one copy of any article(s) in LSE Research Online to facilitate their private study or for non-commercial research. You may not engage in further distribution of the material or use it for any profit-making activities or any commercial gain. You may freely distribute the URL (<http://eprints.lse.ac.uk>) of the LSE Research Online website.

This document is the author's final manuscript version of the journal article, incorporating any revisions agreed during the peer review process. Some differences between this version and the publisher's version remain. You are advised to consult the publisher's version if you wish to cite from it.

<http://eprints.lse.ac.uk>

Contact LSE Research Online at: [Library.Researchonline@lse.ac.uk](mailto:Library.Researchonline@lse.ac.uk)

# **National Carbon Dioxide Emissions:**

## **Geography Matters**

### **REVISED VERSION**

**Dr Eric Neumayer**

Department of Geography and Environment, London School of Economics and Political  
Science, Houghton Street, London WC2A 2AE, UK

Phone: 0207-955-7598. Fax: 0207-955-7412. Email: [e.neumayer@lse.ac.uk](mailto:e.neumayer@lse.ac.uk)

# National Carbon Dioxide Emissions: Geography Matters

*This article examines the role of geographical factors as determinants of cross-country differences in per capita carbon dioxide emissions. Such differences have been explained by economists mostly in terms of per capita income. Geographical factors on the other hand have been neglected by economic analysis. We examine the effects of cold and hot climates, transportation requirements and the availability of renewable energy sources on emissions. We find that with the exception of cooling requirements as measured by hot climates, all these geographical factors are statistically significant determinants of emissions in accordance with our expectation. Furthermore, cold climates and the availability of renewable resources are also substantively important.*

**Key words:** Global, carbon dioxide, quantitative analysis, climate, transportation, renewable energy

## **Introduction**

Economists have analysed the determinants of differences in carbon dioxide (CO<sub>2</sub>) emissions across countries for more than a decade (see, for example, Shafik 1994; Grossman and Krueger 1995; Holtz-Eakin and Selden 1995; Schmalensee et al. 1998; Galeotti and Lanza 1999; Ravallion, Heil and Jalan 2000; Heil and Selden 2001). They mainly explain such differences with the help of per capita income as the independent variable in an analytical framework called the Environmental Kuznets Curve (EKC). In this framework, emissions first rise with increasing income, but at a decreasing rate. For

some pollutants, a turning point is within the range of existing income levels such that emissions are predicted to fall in some countries with high income levels (Cole and Neumayer 2004). In the case of CO<sub>2</sub> emissions, however, the estimated turning point is typically much beyond even the highest existing income level such that emissions merely rise at a decreasing rate with higher income levels, but are not predicted to fall in any country.

Economic analysis has not been interested in the impact of geographical factors on such emissions. Geography should clearly matter, however. For example, we would expect countries faced with cold winter months to have higher heating requirements than others with very mild or even warm winters. Similarly, countries with hot summers might have greater cooling requirements with consequently greater emissions. We would expect countries with populations spatially scattered across a big land area to have higher transportation requirements and therefore higher emissions than small countries or those with highly concentrated clusters of population. Finally, we would expect countries, which nature endowed with the gift of renewable energy resources to have lower emissions. Hydroelectricity, for example, can only be generated if sufficiently voluminous water flows and a suitable topography are available. Wind and solar energy use are most suitable for coast lines and regions with high solar influx. Of course, how a country makes use of its natural endowment of renewable resources is to some extent also determined by political will and economic capacity. But a fundamental dependence on geography exists. Unfortunately, it is not possible to separate one from the other for a global sample.

Economic analysis has not completely overlooked the potential importance of geography, since many studies include country-specific fixed effects or at least model these as random effects forming part of the stochastic error term. Sometimes the

inclusion of country-specific fixed effects is explicitly justified with reference to ‘climatic patterns or resource endowments’ as in Heil and Selden (2001, p. 38). However, the inclusion of such effects is unsatisfactory for mainly two reasons. First, it meshes geographical factors together with any other country-specific fixed effects. Second, in failing to include geographical variables explicitly in the estimations, these studies can offer no insights into the distinct impact of different geographical factors on CO<sub>2</sub> emissions.

This article attempts to demonstrate that geography indeed matters when it comes to CO<sub>2</sub> emissions. It improves upon more preliminary earlier estimates undertaken within a more basic estimation framework (Neumayer 2002). We will show that cold climates, limited access to renewable energy sources and higher transportation requirements are all statistically significantly associated with higher CO<sub>2</sub> emissions. In contrast, cooling requirements as approximated by hot climates exert no statistically significant impact upon emissions.

Somewhat surprisingly, geographical aspects have not played much of a role in the negotiations of emission reductions in Kyoto, Japan, in 1997 (Jones 2001). They have, however, at times been considered in debates about what a fair allocation of emission reduction obligations would look like. For example, Grubb et al. (1992, p. 314) examine, without endorsing, “reasonable emissions” as one criterion for an internationally just allocation rule. They define a ‘reasonable level of emissions for each country’ as the ‘level that would support a consistent, modest standard of living, *given the national climatic and other conditions* [emphasis added]. Permits would be granted for emissions at this level, but not for those “luxury” emissions in excess of this amount’. The Intergovernmental Panel on Climate Change (IPCC 1995, p. 104) contemplates a similar allocation rule under the heading “basic needs”. Such a rule

would allow countries ‘the right to emit the minimum levels of greenhouse gases needed to meet the basic needs of their citizens (...). It would perhaps be close to the allocation of emission permits according to population, although basic needs could vary from country to country *depending on climate and other matters* [emphasis added]’. As a final example, consider the attempt by Benestad (1994) to construct a formula for just allocation of CO<sub>2</sub> emission rights according to energy needs, including such things as a country’s heating and cooling requirements, transportation needs as well as renewable energy sources potential. Since this study examines the relative importance of a number of geographical factors explaining cross-country differences in CO<sub>2</sub> emissions, it can also shed some light on the relevance of these and other normative allocation rules that refer to such factors.

## Research design

We use a panel of per capita CO<sub>2</sub> emissions covering the period 1960 to 1999 with up to 163 countries. The country coverage is entirely driven by the availability of data. We estimate variants of the following model:

$$\ln(E_{it}) = \beta_0 + \beta_1 \ln(Y_{it}) + \beta_2 (\ln Y_{it})^2 + \beta_3 C_i + \beta_4 R_i + \beta_5 A_i + T_t + e_{it}$$

where countries are indicated by  $i$  and years by  $t$ . The variable  $E$  stands for per capita CO<sub>2</sub> emissions,  $Y$  is income per capita,  $C$  is a climate variable,  $R$  is the percentage of total energy consumption derived from renewable energy sources,  $A$  is a measure for transportation requirements,  $T_t$  are year-specific dummy variables and  $e_{it}$  is a stochastic error term. The dependent variable is logged to make its distribution less skewed, which typically makes the estimated model more compatible with distributional assumptions

of the estimation (Wooldridge 2000). The logging of the income variable is common in the EKC literature, partly because it allows an elasticity interpretation of the estimated coefficients and partly because empirically the model fit is better with the income variable logged. Our major results remain valid if these variables were not logged. The year-specific dummy variables capture exogenous advances in carbon saving technology open to all countries.

For panel data, common estimators used are the fixed-effects and either the generalised least squares (GLS) or the full maximum-likelihood random-effects estimators. The fixed-effects estimator cannot be used here since some of our variables do not vary over time. Instead of the GLS or the random-effects estimators we use a generalised estimating equations (GEE) estimator, which is an extension of the generalized linear model (GLM) approach. It is asymptotically equivalent to a random-effects estimator, but has important advantages as well. First, it is easy to compute standard errors that are robust to heteroscedasticity. Second, observations can be assumed to be clustered, which means that they are assumed to be independent only across countries, but not necessarily across time within any one country.<sup>1</sup> Zorn (2001, p. 470) calls for the use of GEE estimators for cases ‘in which the standard assumption that the data are conditionally independent can be called into question’. Since observations within any one country are certainly not independent over time, using standard errors that are robust towards clustering by countries is clearly advantageous.

Total per capita CO<sub>2</sub> emissions from fossil fuel burning and cement manufacturing were taken from data compiled by the Carbon Dioxide Information Center, the definitive source for such data (Marland et al. 2002). Income is measured as real per capita GDP in purchasing power parity taken from the Penn World Table 6.1 (Heston, Summers and Aten 2003). For our climate variables we use two proxy variables for cold

climates and the consequent heating requirements. One is the average minimum temperature in degrees Centigrade in the coldest climatic season of the year. For countries located in the Northern hemisphere this means December, January and February, for those in the Southern hemisphere the temperatures refer to June, July and August. The other variable is the annual number of frost days, which is defined as the number of days in which daily minimum temperature drops below zero degrees Centigrade. Both variables are very highly correlated as one would expect. Neumayer (2002) suggests that the maximum temperature as a proxy for cooling requirements is not consistently and robustly associated with higher emissions. Nevertheless, we use the average maximum temperature in the hottest climatic seasons of the year as a further variable in extended estimations to our main analysis. All variables are taken from the climate data set for political areas described in Mitchell, Hulme and New (2002).

The share of total energy consumption derived from renewable resources is calculated from data in WRI (2003). As this is the variable with the lowest availability, our estimations are run once with and once without it. In addition to hydroelectricity, renewable resources also cover energy from primary solid biomass, thermal solar, photovoltaic solar, wind, biogas, liquid biomass, and tide, wave, and ocean. Fuel and waste renewable energy sources in the form of biomass are much used by poor developing countries. While they partly create CO<sub>2</sub> (and other greenhouse gas) emissions, they are usually not included in CO<sub>2</sub> emission data, which derive exclusively from estimates of fossil fuel burning and cement manufacturing. In as much as fuel and waste renewable energy sources substitute for fossil fuels, which would have otherwise been used, their consumption should lead to lower CO<sub>2</sub> emissions thus measured. In this respect, they do not differ from other substitute renewable energy sources that entail few



CO<sub>2</sub> emissions, such as hydroelectricity. It is therefore correct to include them for the purposes of explaining cross-country CO<sub>2</sub> emissions here.

As concerns transportation requirements, big countries have higher transportation requirements as goods and people are typically moved over longer distances. However, it would be highly misleading to simply take a country's total land area as a proxy for its transportation requirements. This is because often huge parts of big countries such as Canada or the Russian Federation are sparsely inhabited, if at all. To measure transportation requirements, we take two proxy variables. CIESIN (2001) provides geographical information systems data from some time in the late 1990s on the percentage of total land area that is either urbanised (as indicated by lights at night) or used for agriculture. Multiplied by total land area this provides a good proxy to the total land area impacted by human activities and hence for a country's transportation requirements (data for land area taken from World Bank 2002). As an alternative proxy variable, we use the total length of the road network, both paved and unpaved. Poor over-time availability made it necessary to take the average length of the road network in the 1990s, with data taken from IRF (various years). Roads are built where people live. Big countries will have a larger road network, but its length also depends on how scattered the population is over the entire land area. For example, Australia has a small total length of road network relative to its land area as its population tends to be clustered in a few population centres along the coast line. India has a smaller land area, but much larger total length of road network as its population is scattered almost all over the country. Our two proxy variables are very highly correlated as one would hope for given that they are supposed to approximate the same underlying concept transportation requirements. Table 1 provides summary statistics on all variables, table 2 a matrix of bivariate correlation coefficients.

< Insert tables 1 and 2 about here >

## **Results**

We start with the model that excludes the renewable energy variable and therefore has the largest sample size as it. Column I of table 3 shows the expected EKC results: The coefficient of the linear income term is positive and statistically significant, whereas the coefficient of the squared term is negative and significant. A higher minimum temperature during the cold season is associated with lower CO<sub>2</sub> emissions, a longer total road network with higher emissions, all in line with expectations. In column II we replace the temperature variable with the annual number of frost days, which is highly significant with the expected positive sign. In columns III and IV we repeat the first two estimations, but replace the total road length with the variable measuring the land area impacted by human beings. In both estimations, this variable has the expected positive sign, but it is marginally insignificant in column III and only marginally significant at the .1 level in column IV.

< Insert table 3 about here >

In table 4 we similarly estimate four different models analogous to table 3, but add the variable measuring the share of total energy consumption derived from renewable resources. The statistical significance of the existing variables is often slightly reduced in this smaller sample, but all remain significant at the .05 level with the expected signs. The share of renewable resources is negatively associated with emissions and highly

statistically significant. The variable measuring the land area impacted by human beings is now significant in both model specifications at the higher confidence level of .05. This effect is due to the inclusion of the renewable resource share as a further explanatory variable rather than due to the reduction in sample size. This follows from re-running regressions III and IV on the same sample, but without the renewable energy share variable included (detailed results not reported).

< Insert table 4 about here >

We will now include the maximum temperature during the hot climatic season as an additional control variable to our preferred model. Our preferred model is that in which heating requirements are approximated by the annual number of frost days and transportation requirements by the total length of road network. This model is preferred because these variables are more clearly statistically significant than their respective alternatives. We find the maximum temperature variable to be highly insignificant (column I of table 5). This holds true as well for the smaller sample with the renewable resource variable included (column II). Indeed, the same holds true for any other model specification (detailed results not reported).

As a further sensitivity analysis, we tested whether our results are due to the influence of outliers. Belsley, Kuh and Welsch (1980) suggest excluding observations as outliers that have both high residuals and a high leverage. Their criterion is to exclude an observation if its so-called DFITS is greater in absolute terms than twice the square root of  $(k/n)$ , where  $k$  is the number of independent variables and  $n$  the number of observations. DFITS is defined as the square root of  $(h_i/(1-h_i))$ , where  $h_i$  is an observation's leverage, multiplied by its studentized residual. Applying this criterion

leads to the exclusion of 4 countries and 257 observations in the estimation with the larger sample size and the exclusion of 2 countries and 122 observations in the smaller sample where the renewable resource variable is included. Re-estimating our preferred model with the remaining observations leads to the results reported in columns III and IV of table 5. Clearly, the results from our main analysis are not driven by the presence of outliers.

< Insert table 5 about here >

## **Discussion**

In accordance with the existing literature we find a non-linear effect of per capita income levels on per capita CO<sub>2</sub> emissions. Theoretically, therefore, there exists a level of income after which emissions are predicted to decrease with further increases in income. This so-called turning point can be calculated as  $-\beta_1/(2\beta_2)$ , where  $\beta_1$  is the coefficient of the linear and  $\beta_2$  the coefficient of the squared income term. The turning point in our estimations lies between \$55000 as a low estimate (based on results from table 3) and about \$90000 as a high estimate (based on results from table 4). In accordance with the existing literature, we therefore find a turning point that is beyond any currently existing per capita income levels such that CO<sub>2</sub> emissions in all countries are predicted to increase with higher income levels, albeit at a decreasing rate.

What about our geographical variables? To get a feeling for the importance of these variables, it is a good idea to see how much predicted emissions change due to a substantial increase in the variable, where we use a one standard deviation increase to mean substantial. A one standard deviation increase in the average minimum temperature in the cold season reduces CO<sub>2</sub> emissions by between 15 per cent (table 4)

and 41 per cent (table 3).<sup>2</sup> A one standard deviation increase in the annual number of frost days increases emissions by between 22 per cent (table 4) and 71 per cent (table 3). These are clearly non-negligible differences in emissions demonstrating that climatic factors are not only statistically significant, but also substantively important.

In comparison, transportation requirements are less substantively important. The emission increase due to a one standard deviation increase in the total road length is estimated to be between about 8 per cent (table 4) and 17 per cent (table 3). The respective figures for the total land area impacted by human beings are 6 and 9 per cent. These emission increases are clearly more modest compared to the ones for climatic factors.

With respect to the availability of renewable energy sources, a one standard deviation increase in the share of total energy consumption satisfied with renewable sources is predicted to lower CO<sub>2</sub> emissions by about 42 per cent. Again, the availability of renewable energy sources has an impact on CO<sub>2</sub> emissions that is both statistically significant and substantively important.

Another way to gauge the importance of geographical factors is to compare the CO<sub>2</sub> emissions of two fictitious countries. Consider one “average” country at the mean of all independent variables. The second country also has mean income levels, but is geographically disadvantaged in the sense that the annual number of frost days and the total length of road network are one standard deviation above the mean, whereas the renewable resource share is one standard deviation below the mean. Our estimations in column II of table 4 predict that the geographically disadvantaged country has .83 tons higher per capita emissions than the other country. Given that mean CO<sub>2</sub> emissions per capita are .91 tons with a standard deviation of 1.32 tons, this is clearly a non-negligible difference in emissions.

What might explain the insignificance of the maximum temperature variable as a proxy for cooling requirements due to hot climates? One explanation could be that it is a bad proxy variable for cooling requirements. However, there is no strong reason why this should be the case. A perhaps more convincing explanation could be that whereas heating represents a necessity good in cold climates with consumers having few alternatives if they do not want to freeze to death, cooling is likely to be a luxury good in hot climates. Those who can afford will have air conditioning and other cooling devices, those who cannot will not. Many countries with very hot climates are also relatively poor countries with many people not willing or able to afford air conditioning.

On the whole, we have demonstrated that geography matters when it comes to explaining variation in CO<sub>2</sub> emissions. Cold climates and the availability of renewable energy sources exert a statistically significant impact upon such emissions that is also substantively important. We found transportation requirements to be statistically significant as well, but less substantively important.

Our findings become policy relevant when it comes to debates over a fair allocation of emission rights. They clearly show that any simplistic allocation rule on the basis of GDP or population cannot be 'fair' as it would ignore the important role that geography plays in determining cross-country differences in emissions. The results presented above give geographically disadvantaged countries some arguments at hand to request higher emission rights than are given to geographically advantaged countries. The world has only started to embark upon negotiating national emission rights. As emission reduction obligations become tougher in follow-up agreements to the Kyoto Protocol with more countries required to reduce emissions we can expect that countries pay much more attention to geographical factors than they have done so far.<sup>3</sup> Since

geography matters for CO<sub>2</sub> emissions it will also eventually matter for negotiations about emission reductions.

### **Acknowledgements**

I would like to thank two anonymous referees for many helpful comments. Thanks also to Francesca Pozzi from the Center for International Earth Science Information Network (CIESIN), Columbia University, for providing data on one of the variables. I would also like to thank Timothy Mitchell for drawing my attention to their climate data set for political areas.

### **Notes**

<sup>1</sup> At least within Stata, the statistical package used here, there is no easy option to use the random-effects estimators with such robust standard errors. Note that our main results remain valid if we use a random-effects estimator or a GLS estimator with a heteroscedastic and autoregressive error term instead.

<sup>2</sup> Ironically, global warming can be expected to have a small negative feedback effect on carbon dioxide emissions as it is likely to raise minimum temperatures in the cold season as well. According to our estimations, a one degree Centigrade increase in the average minimum temperature would reduce per capita emissions by between 2 (table 4) and 5 percent (table 2).

<sup>3</sup> The same applied to historical accountability for greenhouse gas emissions (Neumayer 2000).

### **References**

**Belsley D A, Kuh E and Welsch R E** 1980 *Regression diagnostics* Wiley, New York

- Benestad O** 1994 Energy needs and CO<sub>2</sub> emissions – constructing a formula for just distributions *Energy Policy* 22 725-734
- CIESIN** 2001 *Data on land area impacted by human activities as a percentage of total land area*. Center for International Earth Science Information Network, Columbia University, New York
- Cole M A and Neumayer E** 2004 Environmental Policy and the Environmental Kuznets Curve in **Dauvergne P** ed *International Handbook of Environmental Politics* Edward Elgar, Cheltenham.
- Galeotti M and Lanza A** 1999 Richer and cleaner? A study on carbon dioxide emissions in developing countries *Energy Policy* 27 565-573
- Grossman G M and Krueger A B** 1995 Economic growth and the environment *Quarterly Journal of Economics* 110 353-377
- Grubb M, Sebenius J, Magalhaes A and Subak S** 1992 Sharing the burden in **Mintzer I M** ed *Confronting climate change: risks, implications and responses* Cambridge University Press, Cambridge 305-322.
- Heil M T and Selden T M** 2001 International trade intensity and carbon emissions: a cross-country econometric analysis *Journal of Environment & Development* 10 35-49
- Heston A, Summers R and Aten B** 2003 *Penn World Table Version 6.1* Center for International Comparisons at the University of Pennsylvania (CICUP)
- Holtz-Eakin D and Selden T M** 1995 Stoking the fires? CO<sub>2</sub> emissions and economic growth *Journal of Public Economics* 57 85-101
- IPCC** 1995 *Climate change 1995 – economic and social dimensions of climate change* Oxford University Press, Oxford
- IRF** various years *World road statistics* International Road Federation, Geneva
- Jones N** 2001 Big chilly countries lose out under Kyoto *New Scientist* 2319 19



- Marland G, Boden T and Andres R J** 2002 *Global, regional and national fossil fuel CO<sub>2</sub> emissions: 1751-1999* Oak Ridge National Laboratory, Carbon Dioxide Information Center
- Mitchell T D, Hulme M and New M** 2002 Climate data for political areas *Area* 34 109-112
- Neumayer E** 2000 In defence of historical accountability for greenhouse gas emissions  
*Ecological Economics* 33 185-192
- Neumayer E** 2002 Can natural factors explain any cross-country differences in carbon dioxide emissions? *Energy Policy* 30 7-12
- Ravallion M, Heil M and Jalan J** 2000 Carbon emissions and income inequality *Oxford Economic Papers* 52 651-669
- Schmalensee R, Stoker, T M and Judson, R A** 1998 World carbon dioxide emissions: 1950-2050 *Review of Economics and Statistics* 80 15-27
- Shafik, N** 1994 Economic development and environmental quality: an econometric analysis  
*Oxford Economic Papers* 46 757-773
- Wooldridge J M** 2000 *Introductory econometrics: a modern approach* South-Western College Publishing, Cincinnati
- World Bank** 2002 *World development indicators* World Bank, Washington DC.
- WRI** 2003 *Earthtrends – the environmental information portal* ([www.wri.org](http://www.wri.org))
- Zorn C J W** 2001 Generalised estimation equation models for correlated data: a review with applications *American Journal of Political Science* 45 470-490

**Table 1**      **Summary statistics**

	<b>N</b>	<b>Mean</b>	<b>Std. Dev.</b>	<b>Min.</b>	<b>Max.</b>
CO <sub>2</sub> p.c. (metric tons)	4688	0.91	1.32	0.01	16.88
GDP p.c. (US\$ of 1996)	4688	5982	6020	281.3	41354
minimum temperature (degree C)	4688	9.7	10.6	-28.1	24.2
maximum temperature (degree C)	4688	29.1	5.7	10.6	43.4
number of frost days (per annum)	4688	39.2	57.5	0	252.5
total length road network (km)	4688	210371	682341	246	6330325
total land area impacted by humans (sq km)	4688	197936	480399	63	2930918
renewable energy (% of energy consumption)	2881	0.31	0.31	0	1

**Table 2 Matrix of bivariate correlation coefficients**

	ln (CO <sub>2</sub> p.c.)	ln (GDP p.c.)	minimum temp.	maximum temp.	frost days	length road network	land area impacted
ln (GDP p.c.)	.8972						
minimum temperature	-.5530	-.5295					
maximum temperature	-.5051	-.5788	.7134				
frost days	.5407	.5349	-.9269	-.8338			
length road network	.2034	.1755	-.2217	-.0351	.2060		
land area impacted	.1026	.0234	-.2264	-.0096	.2004	.8282	
renewable energy share	-.8399	-.7462	.5179	.3058	-.4028	-.1457	-.0657

Note: N = 2881.

**Table 3 Results with renewable energy share excluded**

	I	II	III	IV
ln (GDP p.c.)	2.72***	2.72***	2.71***	2.71***
	(3.99)	(3.98)	(3.98)	(3.97)
[ln (GDP p.c.)] <sup>2</sup>	-.12***	-.12***	-.12***	-.12***
	(3.04)	(3.05)	(3.03)	(3.04)
minimum temperature	-.05***		-.05***	
	(6.83)		(6.53)	
frost days		.01***		.01***
		(6.70)		(6.45)
length road network	2.22e-07***	2.26e-07***		
	(3.41)	(3.07)		
land area impacted			1.63e-07	1.80e-07*
			(1.51)	(1.71)
# countries	163	163	163	163
# observations	4688	4688	4688	4688

Note: Dependent variable is ln (CO<sub>2</sub> p.c.). General Estimation Equations (GEE) estimation.

Absolute z-values in parentheses. Robust standard errors allowing observations to be clustered by countries. Coefficients of constant and year-specific time dummies not shown.

\* statistically significant at .1 level \*\* at .05 level \*\*\* at .01 level.

**Table 4 Results with renewable energy share included**

	<b>I</b>	<b>II</b>	<b>III</b>	<b>IV</b>
ln (GDP p.c.)	2.50***	2.48***	2.48***	2.46***
	(3.01)	(3.02)	(2.99)	(3.00)
[ln (GDP p.c.)] <sup>2</sup>	-.11**	-.11**	-.11**	-.11**
	(2.32)	(2.33)	(2.30)	(2.30)
renewable energy share	-1.77***	-1.79***	-1.78***	-1.80***
	(6.03)	(6.30)	(6.01)	(6.29)
minimum temperature	-.02***		-.02**	
	(2.60)		(2.41)	
frost days		.003***		.003***
		(3.33)		(3.15)
length road network	1.23e-07***	1.12e-07***		
	(3.81)	(3.84)		
land area impacted			1.38e-07**	1.20e-07**
			(1.98)	(1.94)
# countries	119	119	119	119
# observations	2881	2881	2881	2881

Note: Dependent variable is ln (CO<sub>2</sub> p.c.). General Estimation Equations (GEE) estimation.

Absolute z-values in parentheses. Robust standard errors allowing observations to be clustered by countries. Coefficients of constant and year-specific time dummies not shown.

\*\* statistically significant at .05 level \*\*\* at .01 level.

**Table 5**      **Sensitivity analysis**

	<b>I</b>	<b>II</b>	<b>III</b>	<b>IV</b>
ln (GDP p.c.)	2.71***	2.51***	3.32***	2.12***
	(3.96)	(3.03)	(5.66)	(3.37)
[ln (GDP p.c.)] <sup>2</sup>	-.12***	-.11**	-.15***	-.09**
	(3.03)	(2.34)	(4.29)	(2.47)
renewable energy share		-1.75***		-1.84***
		(6.06)		(7.35)
frost days	.012***	.005***	.009***	.003***
	(6.30)	(3.45)	(7.69)	(3.94)
length road network	2.05e-07***	9.85e-08***	1.86e-07***	1.02e-07***
	(2.58)	(2.93)	(3.14)	(4.12)
maximum temperature	.03	.02		
	(1.23)	(1.14)		
# countries	163	119	159	117
# observations	4688	2881	4431	2759

Note: Dependent variable is ln (CO<sub>2</sub> p.c.). General Estimation Equations (GEE) estimation.

Absolute z-values in parentheses. Robust standard errors allowing observations to be clustered by countries. Coefficients of constant and year-specific time dummies not shown.

\*\* statistically significant at .05 level \*\*\* at .01 level.