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Discussion of J.F.Bjørnstad, 'Non-Bayesian multiple imputation'

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Non-Bayesian Multiple Imputation by Jan Bjørnstad

Comments by Chris Skinner, University of Southampton

Two main types of application of multiple imputation (MI) to official statistics have been proposed:

1. the use of MI by a statistical agency in its ‘primary’ activity of producing estimates from survey data;
2. the construction of multiply imputed datasets for ‘secondary’ analysis by different users, including the construction of synthetic datasets to protect confidentiality (Raghunathan et al., 2003).

Jan Bjørnstad’s paper focuses on the first use and I shall restrict my comments to this topic also. See e.g. Meng (1994) on some of the issues arising with the second use.

The condition for MI to lead to valid inference in a non-Bayesian framework is termed ‘proper’ by Rubin (1987, Ch. 4). I think it is important to recognize that this condition applies not to the imputation method alone, but to the imputation method for a given ‘complete data’ point estimator and variance estimator (i.e. $\hat{\theta}$ and $\hat{V}(y)$ in the notation here). Thus an agency could not necessarily make the standard MI approach valid by drawing imputations in a Bayesian way from the predictive distribution of the missing values, even if this were practical, since such imputation might not be proper with respect to the agency’s methods of point and variance estimation, and thus might lead to biased variance estimation (Wang and Robins, 1998; Nielsen, 2003; Kim et al., 2006). Instead, the agency would either need to determine an imputation procedure which

was proper with respect to its estimation methods, which Binder and Sun (1996) have shown is often extremely difficult for the kinds of methods used by agencies in practice, or the agency would have to consider revising its (complete data) point and variance estimation procedures to fit in with the imputation procedure, which might be viewed as the ‘tail wagging the dog’.

Jan Bjørnstad’s paper considers to what extent the validity of MI might be retained through the less extreme option of modifying the MI variance estimator, while retaining the agency’s preferred (complete data) point and variance estimator and imputation method. While one could consider the problem of estimating the variance of a MI point estimator for a non-proper imputation scheme from first principles, as in Wang and Robins (1998), Jan Bjørnstad explores instead the possibility of making a simple modification to the standard MI combination formula via the use of the k term in expression (1).

The paper provides interesting demonstrations, for a number of specific cases, that consistent variance estimation can be achieved by taking k to be a measure of the proportion of missing information, such as the reciprocal of an item response rate. For the approach to provide a principled basis for general applications, in line with the aims of MI, it is desirable to understand the method’s potential generality, as discussed in Section 6 and summarised in the Theorem. Two basic conditions of the Theorem seem to me reasonably uncontentious. A number of authors have considered combinations of estimators and imputation methods which obey $E(\hat{\theta}^* | y, s) = \hat{\theta}$ (c.f. Rubin, 1987, equation 4.2.5; Binder and Sun, 1996, equation 14; Kim et al., 2006, condition C3) to restrict attention to cases where the MI point estimator is unbiased. Likewise, condition (3)

corresponds to common assumptions (c.f. Rubin, 1987, equation 4.2.8; Binder and Sun, 1996, equation 19; Kim et al., 2006, equation 3.4).

However, in a number of other respects, the conditions (a)-(g) used in A.5 of the Appendix to prove the Theorem seems rather restrictive. Firstly, the MCAR condition (or conditions a and b) has very strong consequences, especially by preventing consideration of estimators with unequal weights (see the Lemma in section 6) or domain estimators, whereas Kim et al. (2006) show that these are particularly important features of official statistics applications which may lead to bias in the multiple imputation variance estimator. Further restrictions in the Appendix are that: the result seems restricted to a limited class of sampling schemes, for example multi-stage sampling is not discussed; to a limited range of imputation schemes, for example excluding most versions of nearest neighbour imputation methods widely used in official statistics, and to statistics with imputed values in only one variable.

There seem to me therefore to be a number of dimensions of generality that would be useful to research further. Moreover, to assess the extent to which the proposed approach may be useful in practice, it seems necessary to assess the relative merits of the proposed approach with alternative methods for variance estimation with imputed data such as by Rao and Shao (1992), Shao and Steel (1999) and Kim and Fuller (2004). Criteria for comparison include the breadth of conditions under which the approaches are valid and the extent to which the methods provide unified approaches for sets of estimands, such as means, totals or proportions, across different domains and different variables. Other standard criteria are efficiency of point and variance estimation and ease of computation.

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