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Themed Section: The Health Economics of Alzheimer's Disease and Related Dementias

IPECAD Modeling Workshop 2023 Cross-Comparison Challenge on Cost-Effectiveness Models in Alzheimer's Disease

Ron Handels, PhD, William L. Herring, PhD, Farzam Kamgar, MSc, Sandar Aye, MSc, Ashley Tate, PhD, Colin Green, PhD, Anders Gustavsson, PhD, Anders Wimo, MD, PhD, Bengt Winblad, MD, PhD, Anders Sköldunger, PhD, Lars Lau Raket, PhD, Chelsea Bedrejo Stellick, MSc, Eldon Spackman, PhD, Jakub Hlávka, PhD, Yifan Wei, PhD, Javier Mar, PhD, Myriam Soto-Gordoa, PhD, Inge de Kok, PhD, Chiara Brück, PhD, Robert Anderson, MA, Peter Pemberton-Ross, PhD, Michael Urbich, PhD, Linus Jönsson, MD, PhD

ABSTRACT

Objectives: Decision-analytic models assessing the value of emerging Alzheimer's disease (AD) treatments are challenged by limited evidence on short-term trial outcomes and uncertainty in extrapolating long-term patient-relevant outcomes. To improve understanding and foster transparency and credibility in modeling methods, we cross-compared AD decision models in a hypothetical context of disease-modifying treatment for mild cognitive impairment (MCI) due to AD.

Methods: A benchmark scenario (US setting) was used with target population MCI due to AD and a set of synthetically generated hypothetical trial efficacy estimates. Treatment costs were excluded. Model predictions (10-year horizon) were assessed and discussed during a 2-day workshop.

Results: Nine modeling groups provided model predictions. Implementation of treatment effectiveness varied across models based on trial efficacy outcome selection (clinical dementia rating – sum of boxes, clinical dementia rating – global, mini-mental state examination, functional activities questionnaire) and analysis method (observed severity transitions, change from baseline, progression hazard ratio, or calibration to these). Predicted mean time in MCI ranged from 2.6 to 5.2 years for control strategy and from 0.1 to 1.0 years for difference between intervention and control strategies. Predicted quality-adjusted life-year gains ranged from 0.0 to 0.6 and incremental costs (excluding treatment costs) from –US\$66 897 to US\$11 896.

Conclusions: Trial data can be implemented in different ways across health-economic models leading to large variation in model predictions. We recommend (1) addressing the choice of outcome measure and treatment effectiveness assumptions in sensitivity analysis, (2) a standardized reporting table for model predictions, and (3) exploring the use of registries for future AD treatments measuring long-term disease progression to reduce uncertainty of extrapolating short-term trial results by health-economic models.

Keywords: Alzheimer's disease, cross-validation, decision-analytic modeling, health-economic evaluation.

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Introduction

Health-economic studies grounded in mathematical models play a pivotal role in supporting policy discussions concerning the accessibility, reimbursement, and pricing of interventions for Alzheimer's disease (AD). These model-based evaluations rely on results from clinical trials, necessitating the extrapolation of short-term clinical efficacy (eg, effects on biomarker targets and cognitive scales) to estimate the lifetime impact on patient-relevant outcomes (eg, function, care needs, and mortality). This in turn is often based on evidence from disease registries and cohort studies. Different healtheconomic models have shown to produce dissimilar results. Hence, transparency and credibility of these models are key to decision

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makers who rely on them¹ to facilitate timely access to new interventions for people living with AD. This article describes research aimed at better understanding why these differences

among model predictions exist.

Comparative modeling has the potential to increase confidence in models if similar model predictions are observed.^{2,3} Unexplained differences across AD model predictions could jeopardize

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Highlights

- The results of our cross-comparison study suggest that the following cascade of factors explained much of the variation in Alzheimer's disease (AD) health-economic model predictions: model design drove the choice of trial outcome, which determined the relative effect size, which together with natural history and waning/ discontinuation assumptions drove the gained time in mild cognitive impairment, which drove qualityadjusted life-years gains and costs.
- For the evaluation of new AD treatments, we recommend (1) addressing trial outcome choice and treatment effectiveness assumptions in sensitivity analysis, (2) the use of a standardized reporting table for model predictions to support model cross-comparison, and (3) exploring the use of registries to measure long-term disease progression of future AD treatments to reduce the uncertainty of extrapolating short-term trial results by health-economic models.
- By identifying possible causes for variation between AD model predictions, our study supports AD model transparency and credibility and generates a set of recommendations for modelers and policy makers.

their credibility, making them unusable to support decision makers (eg, as reported in the field of abdominal aortic aneurysm screening^{4,5}).

In 2020, the International Pharmaco-Economic Collaboration on Alzheimer's Disease (IPECAD) hosted a workshop to crosscompare AD models. Modeling groups were asked to implement a benchmark scenario consisting of a treatment with its effectiveness conceptually defined as a "30% reduction in disease progression."⁶ Models were compared on the health-economic predictions they produced, which resulted in an increased understanding of model differences. However, because the benchmark treatment effect was defined on a conceptual level (ie. 30%). modelers did not have challenges on methodological (premodel analysis) choices on implementing real-world short-term clinical trial efficacy outcomes into their model, as well as extrapolation of effectiveness beyond trial follow-up evidence to long-term healtheconomic estimates. This challenge was discussed among modelers during the 2021 IPECAD follow-up workshop, highlighting the disconnect between efficacy estimates reported in AD clinical trials and effectiveness inputs required for their health-economic models. More specifically, challenges included (1) discrepancies in endpoints and scales in trials compared with models, (2) treatment effects reflecting single or multiple domains, (3) differences in statistical methods for analyzing trial efficacy data, and (4) differing analytical objectives for trials (identify a treatment effect in a clinical setting) compared with models (estimate the treatment benefit in a routine care setting) (see Appendix 1 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2 024.09.006 for details).⁷

Given our previous workshop recommendations for aligning models, the raised challenges around implementing a treatment effect, the recent Food and Drug Administration approvals of new AD treatments,^{8,9} and the recently published significant results for a similar drug,¹⁰ we believe there remains an urgent need to cross-compare AD models to increase our understanding of them and to support their credibility for evaluating new AD treatments.

This initiative aimed to cross-compare predictions from decision-analytic models for AD that start in mild cognitive impairment (MCI) and have implemented the same hypothetical trial efficacy estimates. The following research question was the starting point: What are the differences in key health-economic predictions across models that assess the cost-effectiveness of a hypothetical disease-modifying AD drug treatment, and what factors related to the use of clinical trial evidence explain those differences?

This article describes the methodology and results from the 2023 IPECAD workshop on the cross-comparison of AD models. The focus of this cross-comparison was on the challenges of implementing realistic short-term trial efficacy estimates in a decision model and extrapolating them to long-term health-economic estimates. Our goal was to describe treatment effect implementation methods and discuss how they translated into differences in health-economic model predictions, taking into account the specific model design, parameterization, and model assumptions.

Methods

The IPECAD steering group (ie, organizers) a priori defined a research question based on the previous workshop recommendations⁶ (see introduction). Then, they designed a model cross-comparison study based on guidelines for multi-model comparisons¹¹ and foundations for model validation and comparison.² It consisted of the following 5 main steps.

Second, the organizers drafted a benchmark scenario, shared it with participating modeling groups, and provided them with the possibility to provide feedback to refine it.

Third, each modeling group implemented the benchmark scenario independently into their model as they deemed appropriate. Modeling groups were asked to maintain their model structures to avoid model convergence (ie, harmonization on [almost] all parameters omitting genuine discrepancies between models). Modeling groups were asked to report details on how they implemented the benchmark scenario into their model and to report their model predictions (ie, time in AD disease states, costs, and quality-adjusted life-years [QALYs]) to the organizers.

Fourth, organizers validated, cross-compared, and summarized all implementation details and model predictions and shared them with all modeling groups.

Fifth, we discussed cross-compared implementation details and model predictions during a workshop held on June 5 and 6, 2023, in Stockholm.

Model Identification, Selection, and Invitation (Step 1)

Models were identified through earlier participation in IPECAD workshops, systematic reviews, a PubMed search, an open call through ISPOR, snowballing, and ad hoc identification by the organizers. Duplicate models were omitted, leaving 46 identified models.

Modeling groups were invited (see Appendix 2 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2 024.09.006) if their models fulfilled the inclusion and exclusion criteria (reflected AD starting in MCI, AD disease progression, and health-economic outcomes in a US clinical setting) or could potentially fulfill the criteria after relatively small adjustments (n = 38). Among those invited, 12 did not respond, 2 considered their model outdated, 10 indicated insufficient resources to implement the benchmark scenario, and 4 did not have an operational model. This left 10 modeling groups for participation in the workshop (see Appendix 3 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2024.09.006 for details).

Benchmark Scenario and Inputs (Step 2)

A hypothetical 18-month placebo-controlled trial for an AD treatment was developed in terms of eligibility criteria, baseline characteristics, and trial outcome estimates reflecting a population of persons with a diagnosis of MCI due to AD set in a US memory clinic setting. This was operationalized by creating summary tables and figures from synthetic individual patient-level data¹² on surrogate and secondary outcome estimates using observational data from the Alzheimer's Disease Neuro-imaging Initiative¹³ (see Appendix 4 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2024.09.006).

The benchmark scenario to be implemented in the model was based on previous workshop recommendations⁶ as follows. The target population was set as a standardized patient aged 70 years, male or female (modeled separately), with a diagnosis of amyloidconfirmed AD-type MCI on their first visit to a memory clinic in the United States. To focus on treatment implementation only, costs and effects related to diagnostics fell outside the scenario. The intervention strategy was a hypothetical treatment in addition to standard of care (ie, as reflected in the hypothetical 18-month trial results). The control strategy was standard of care. The treatment effect was considered disease modifying, although the implications of disease modification in a modeling context were not specified. Treatment discontinuation was set at 10% of patients

Table 1. Implementation of population and setting and natural history progression in MCI.

| Group | Markov type | | | | | | Microsimulation type | | | |
|--|---|---|---|--|---|---|--|--|---|--|
| Participant | Biogen | CPEC | IPECAD | Bedrejo and colleagues | SveDem | BASQDEM | FEM | RTI-HS | MISCAN- Dementia | |
| Model information | | | | | | | | | | |
| Abbreviation | (company name) | Care Policy Evaluation Centre | International Pharmaco- Economic Collaboration on Alzheimer's Disease | (author name) | Swedish Dementia Registry | Basque dementia model | Future Elderly Model* | (company name) | Microsimulation Screening Analysis (Dementia) | |
| Developer(s) | P.PR. [†] M.U. [†] (Biogen), W.L.H. [†] (RTI- HS) | R.A. [†] (London School of Economics), R.H. [†] (Maastricht University) | R.H. [†] (Maastricht University), C.G. [†] (Biogen), A.G. [†] (Quantify Research) | C.B.S. [†] (University of Alberta), E.S. (University of Calgary) | A.W., R.H. [†] (Karolinska Institute) | J.M. [†] (Osakidetza Basque Health Service), M.S G. [†] (Mondragon Unibertsitatea) | B.T. (USC), Y.M. [†] (USC), J.H. [†] (USC, Masaryk University) | W.L.H., [†] F.K. [†] (RTI Health Solutions) | C.B., [†] l.d.K. [†] (Erasmus Medical Center) | |
| Reference | Herring et al ¹⁴ | Anderson et al ¹⁵ | Green et al ¹⁶ | | Wimo et al ¹⁷ | Mar et al ¹⁸ | Goldman et al ¹⁹ | Herring et al ²⁰ ; additional details in Handels et al ^{6,21} | Bruck et al ²² | |
| Open source | No | No | Yes v1.1.2 www.ipecad. org | No | Yes v2.2.0 https://gi thub.com/ronhande ls/IPECAD | No | Yes: https:// hrsdata.isr. umich.edu/ data-products/ future-elderly- model-fem-files [‡] | No | No | |
| MCI natural history | | | | | | | | | | |
| Data source(s) | Synthetic trial control arm | Vos et al ²³ | Vos et al ²³ | Potashman et al ⁴⁰ [2021] | Synthetic trial control arm | Synthetic trial control arm; van Oudenhoven et al ⁴⁴ [2019] | Cognitive status of US population in the Health and Retirement Study (HRS); Wei et al ⁴⁶ (2022) | Synthetic trial control arm; Sapkota et al ⁴² [2021] | Synthetic trial control arm; Vermunt et al ⁴⁵ [2019] | |
| Source's estimate (and implementation) | Annualized CDR-global transition probability from CDR- global 0.5 (reflecting MCI with A β +) to mild AD with CDR-global 1; backward transitions (CDR-global 0.5 to CDR- global 0) to no AD assumed to remain in MCI | Annualized transition probability for IWG-2 prodromal AD criteria (reflecting MCI with Aβ+) to AD dementia | Annualized transition probability based on weighted average of NIA-AA high AD likelihood group and NIA-AA conflicting IAP group (reflecting MCI with Aβ+) to AD dementia | Annual CDR-SB transition probability from CDR-SB 0.5-4.0 (reflecting MCI with $A\beta$ +) to mild AD with CDR-SB 4.5-9; backward transitions to asymptomatic AD assumed to remain in MCI | Annualized CDR- global transition probability from CDR- global 0.5 (reflecting MCI with $A\beta$ +) to mild AD with CDR-global 1; backward transitions (CDR-global 0.5 to CDR-global 0) to no AD assumed to remain in MCI | Mixed model (originally developed based on van Oudenhoven et al ⁴⁴ [2019] CDR-SB data; differentiation between moderate and severe added to model. Model simulates times until mild, moderate, and severe using cut points CDR-SB 4.5, 9.5, and 16, respectively. | ADNI: linear regression model for CDR- SB using splines of lagged CDR- SB, age, gender, and race as covariates. Synthetic trial: baseline CDR-SB and gender (separate) used as input for linear model and cutoff applied to classify as mild, moderate, and severe dementia | Sapkota: annual decline in MMSE and DAD (reflecting MCI with $A\beta+$) using MMSE and FAQ changes from Synthetic trial (FAQ mapped to DAD) Synthetic trial: baseline MMSE and DAD (mapped from FAQ) | Vermunt: duration from MCI with Aβ+ to AD dementia, calibrated to proportion transitions to CDR-global 1 (from synthetic trial) by restricting fast progressing individuals; backward transitions (CDR 0.5 to CDR 0) to no AD not considered | |
| Source's age and sex (and adjustments) | Age: 73 years, gender: 42% female | Age: 71.4 years, gender: 48% female) | Age: 71.4 and 66.1 years, gender: 47% and 43% female, respectively | Age and gender for MCI patients not reported separately | Age: 73 years, gender: 42% female | Age: 73 years, gender: 42% female | Age: 70 years (starting age in model) Gender: gender- specific progression | Age: 73 years, gender: 42% female | Age: 70 years, gender: 42% (reflected as observed in data source), with gender-specific progression in the model calibrated to gender-average rate observed from trial | |

Table 1. Continued

| Group | Markov type | | | | | | Microsimulation type | | | |
|---|--|--|---|--|--|--|---|---|---|--|
| Participant | Biogen | CPEC | IPECAD | Bedrejo and colleagues | SveDem | BASQDEM | FEM | RTI-HS | MISCAN- Dementia | |
| Source's syndrome(s) (and adjustments) | Prevalent MCI defined by episodic memory impairment and CDR- global (0.5) and MMSE (≥24) at baseline; conversion to dementia defined by CDR-global score | Prevalent MCI based on Petersen et al ⁴⁹ criteria; conversion to AD-type dementia defined by clinical criteria | Prevalent MCI based on Petersen et al (2004) criteria; conversion to AD-type dementia defined by clinical criteria | Incident MCI defined by CDR- SB (0.5-4.0) at baseline; conversion to dementia defined by CDR- SOB | Prevalent MCI defined by episodic memory impairment and CDR-global (0.5) and MMSE (≥24) at baseline; conversion to dementia defined by CDR-global score | Prevalent MCI defined by episodic memory impairment and CDR- global (0.5) and MMSE (≥24) at baseline; conversion to dementia defined by CDR-global score | Prevalent MCI defined by CDR- global (0.5) at baseline; conversion to dementia defined by CDR- global (≥1) | Prevalent MCI defined by CDR- global (0.5) and MMSE (\geq 24) at baseline; AD dementia diagnosis at MMSE \geq 1.5 SDs below normal and DAD \leq 93 applied to simulated data (informed by MCKhann et al ³⁷ [2011]) | Prevalent MCI defined by episodic memory impairment and CDR-global (0.5) and MMSE (≥24) at baseline; conversion to dementia defined by CDR- global score | |
| Source's pathology (and adjustments) | Amyloid positive via abnormal CSF amyloid beta or amyloid PET; no criteria on tau | IWG-2 criteria: "Abnormal CSF amyloid- beta1-42 and CSF tau" or "abnormal amyloid PET" | NIA-AA high criteria: Abnormal amyloid (CSF or PET) and abnormal neuronal injury markers (CSF, PET, or MRI); NIA-AA conflicting IAP criteria: abnormal amyloid (CSF or PET) and normal neuronal injury markers (CSF, PET, or MRI) | Amyloid positive via abnormal CSF amyloid beta, amyloid PET, or (retrospectively) autopsy result; no criteria on tau or neuronal injury | Amyloid positive via abnormal CSF amyloid beta or amyloid PET; no criteria on tau | Amyloid positive via abnormal CSF amyloid beta or amyloid PET; no criteria on tau | No amyloid parameter or assumption has been made | Synthetic trial: Amyloid positive via abnormal CSF amyloid beta or amyloid PET; no criteria on tau | Synthetic trial: Amyloid positive via abnormal CSF amyloid beta or amyloid PET; no criteria on tau | |
| Source's clinical setting (and adjustments) | Clinical trial | Combination of local, academic, and tertiary memory clinic, and research care settings | Combination of local, academic, and tertiary memory clinic, and research care settings | Research care settings | Clinical trial | Clinical trial | ADNI: combination of local, academic, and tertiary memory clinic and research care settings | Clinical trial | Clinical trial | |
| Model severity level cutoffs | CDR-SB, MCI: 0.5-4, mild: 4.5-9, moderate: 9.5-15.5, severe: 16- 18 [O'Bryant et al, ³⁹ 2010] | MMSE, mild: 30-20, moderate: 19- 10, severe: 9- 0 | MMSE, mild: 30-20, moderate: 19- 10, severe: 9- 0 FAQ, mild: 0- 8, moderate: 9-23, severe: 24-30) NPI-Q, mild: each item ≤ 1 , moderate: each item ≤ 1 , moderate: each item ≤ 2 with at least one 2, severe: at least one item with 3 | CDR-SB, MCI: 0.5-4, mild: 4.5- 9, moderate: 9.5-15.5, severe: 16-18 [O'Bryant et al, ³⁹ 2010] | MMSE, mild: 30-20, moderate: 19-10, severe: 9-0 | CDR-SB, MCI: ≤4, mild: 4.5-9, moderate: 9.5- 15.5, severe: 16-18 [O'Bryant et al, ⁵⁹ 2010] | Crosswalk [O'Bryant et al, ³⁹ 2008: table 2] to generate global CDR from simulated CDR- SB | MMSE, mild: >20, moderate: 10-20, severe: <10 | CDR-global, MCI: 0.5, mild: 1, moderate: 2, severe: 3 | |

continued on next page

Table 1. Continued

| Group | Markov typ | be | | | | Microsimula | tion type | | |
|--|--|--|---|--|--|--|--|---|---|
| Participant | Biogen | CPEC | IPECAD | Bedrejo and colleagues | SveDem | BASQDEM | FEM | RTI-HS | MISCAN- Dementia |
| Model settings for the United States | | | | | | | | | |
| Costs | Direct medical costs: Leibson et al ³⁵ [2015] Direct nonmedical costs: Robinson et al ⁴¹ [2020]; Gustavsson et al ³⁰ [2011] | Gustavsson et al ³⁰ [2011] and assumed to be over and above the age- specific nondementia cost of health services. MCI was assumed to not impose excess cost. | Dementia: Gustavsson et al ³⁰ [2011]; MCI: assumed 50% of mild dementia [Darba et al, ²⁷ 2015; Jonsson et al, ³¹ 2006] | Lin et al ³⁶ [2021] ICER report (derived from Leibson et al ³⁵ [2015]) | Tahami et al ⁴³ [2022] (derived from Robinson et al ⁴¹ [2020] and Genworth et al ²³ [2022]) weighted for institutionalization proportions (derived from Tahami Monfared et al ²⁴ referring to Neumann et al ³⁸ [1999]). | Robinson et al ⁴¹ [2020]; Gustavsson et al ³⁰ [2011] | Medicare Current Beneficiary Survey (2007- 2010) | Robinson et al ⁴¹ [2020]; Gustavsson et al ³⁰ [2011] | Robinson et al ⁴¹ [2020]; Gustavsson et al ³⁰ [2011] |
| Utilities | Landeiro et al ³³ [2020] | Neumann et al ³⁸ [1999]; Karlawish et al ³² [2008] | Neumann et al ³⁸ [1999] | Neumann et al ³⁸ [1999] | Neumann et al ³⁸ [1999] | Neumann et al ³⁸ [1999] | Leaf ³⁴ [2015] | Landeiro et al ³³ [2020] | Neumann et al ³⁸ [1999] |
| Mortality | RRs for MCI and dementia [Wilson et al, ⁴⁷ 2009] applied to US general population mortality 2019 | Excess mortality in moderate and severe dementia Brookmeyer et al ²⁶ [2007] applied to US general population mortality 2019 | RRs for dementia only [Andersen et al, ²⁵ 2010] applied to US general population mortality 2019 | RRs for MCI and dementia [Andersen et al, ²⁵ 2010] applied to US general population mortality 2020 | RRs for dementia only ¹⁷ relative to very mild dementia and Andersen et al ²⁵ [2010] very mild dementia to normal) applied to US general population mortality 2019 | Time to death from mild- moderate- severe dementia [Dodge et al, ²⁸ 2003]; time to death from other causes adjusted to US setting from US life table | Mortality risk matches those observed in HRS adjusted for risk factors. Data from the 2018 wave of HRS are used | RRs for MCI and dementia ¹⁷ (see also Wilson et al ⁴⁷ 2009) applied to US general population mortality 2019 | Time to death from dementia from Rotterdam study [Wolters et al, ⁴⁸ 2019]; time to death from other causes from US general population life tables 2019 |

A β + indicates amyloid beta positive; AD, Alzheimer's Disease; ADNI, Alzheimer's Disease Neuroimaging Initiative; CDR, clinical dementia rating; CDR-SB, clinical dementia rating - sum of boxes; CSF, cerebrospinal fluid; DAD, disability assessment for dementia; FAQ; functional activities questionnaire; FEM, future elderly model; HRS, Health and Retirement Study; IAP, isolated amyloid pathology; MCI, mild cognitive impairment; MMSE, mini-mental state examination; MRI, magnetic resonance imaging; NIA-AA, National Institute on Aging and the Alzheimer's Association; PET, positron emission tomography; RR, relative risk; SD, standard deviation. *FEM draws on a cognitive and functional impairment transition model that has been validated for a 5-year period. Long-term estimates introduce considerable

Increasing and may underestimate or overstimate the value of treatment in the scenario studied. [†]Indicates present during the workshop.

[‡]Details can be found here: https://www.rand.org/content/dam/rand/pubs/rgs_dissertations/RGSDA1400/RGSDA1439-1/RAND_RGSDA1439-1.pdf; please contact authors for details on methods when replicating the results of this study.

over the 18-month trial follow-up period, which included discontinuation due to any adverse events (including amyloid-related imaging abnormalities). Treatment was to be applied up to and including mild dementia. Treatment waning was set at 5% per year. US-specific background mortality and consumer price index data were specified for the year 2019 (reflecting pre-COVID-19 levels). The time horizon was set at 25 years and the discount rate at 3.5% per year for both costs and health outcomes. Suggestions to reflect a US setting in terms of mortality, utility, and cost estimates were also provided (see Appendix 5 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2 024.09.006).

Implement Benchmark Scenario and Model Predictions (Step 3)

Both the synthetic individual-patient level data set and summary tables and figures were provided to all modeling groups. If applicable, modeling groups were expected to apply mapping to translate the benchmark scenario outcomes to the outcomes used in their model.

Each modeling group was asked to implement the benchmark scenario. Model predictions consisted of proportions in AD disease states over time, cumulative time in AD disease states, and healtheconomic outcomes (QALYs and costs) by control/intervention strategy and sex. These model predictions were requested to be shared in a predefined format (see Tables 1^{14-48} , 2, and 3).

Validation and Analysis (Step 4)

The participating modeling groups submitted their benchmark scenario implementation details and model predictions, and the organizers synthesized them into summary tables and graphs. An ad hoc quality check was performed to compare the person-years alive with the person-years in disease states. A gualitative review of the implementation description was also performed. If issues were identified, either from these consistency checks or during the workshop discussion, modeling groups had the opportunity to revisit their approach and submit updated results. This resulted in the following changes after the workshop: (1) adjusting MCI progression to better reflect the changes from baseline over time observed in the synthetic trial data (RTI-HS model), (2) use of an updated open-source version for reasons unrelated to the results (SveDem model), (3) redesign of the benchmark implementation due to incorrect interpretation (Bedrejo model), and (4) omitting the results due to lack of face validity (ADDITION model⁵⁰). A member of the future elderly model (FEM) group participated in the workshop discussions without having submitted their model predictions due to time limitations. We relied on any validation and verification of the original models, and we similarly relied on

the individual modeling groups to validate their implementations of the benchmark scenario. Modeling groups did not have access to one another's model predictions before submitting their results, except for the FEM, although the use of other model predictions for preparing their own model predictions was viewed as unlikely.

Workshop and Results Synthesis (Step 5)

All results (implementation details, model predictions, and summarized cross-comparison results) were shared with all modeling groups before the workshop. Next, participants presented and discussed these results (except for FEM results) during the workshop in a hybrid setting, with online participants' discussion occurring separately among themselves and summarized to the face-to-face participants. Given that there was no established methodology for evaluating concordance in healtheconomic simulation modeling, between-model differences were interpreted qualitatively.

Results

In this section, we summarize the treatment implementation methods and the variability between model predictions. In addition, we describe possible causes for variation based on discussions during the workshop and on supporting post hoc correlation analyses provided in Appendix 6 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2024.09.006.

Treatment Implementation Methods

The following 9 participating models (5 Markov and 4 microsimulation type) were part of this cross-comparison: BASQDEM,¹⁸ Biogen,¹⁴ CPEC,¹⁵ FEM,¹⁹ RTI-HS,²⁰ IPECAD open-source model,¹⁶ MISCAN,²² Bedrejo and colleagues, and SveDem open-source model¹⁷ (Table 1).¹⁴⁻⁴⁸

The sources used to represent the target population's natural disease progression varied in setting (clinical vs general population) and biomarker status (abnormal amyloid only, both abnormal amyloid and neuronal injury, implied amyloid abnormality). Some models modified their MCI natural history to fully (Biogen, SveDem) or in part (BASQDEM, RTI-HS, MIS-CAN) reflect the synthetic trial placebo arm. The specific starting age of 70 years and separately modeled sex were mainly applied to mortality whereas the average age and sex of the source for natural history were assumed representative for the modeled population (except for MISCAN, BASQDEM, FEM). See Table 1¹⁴⁻⁴⁸ for details.

The implementation of the treatment effect varied in choice of outcome domain (cognition, function, and/or composite), choice of scale (clinical dementia rating - sum of boxes [CDR-SB], CDR-global, mini-mental state examination [MMSE], and/or functional activities questionnaire [FAQ]), and choice of unit (health state transition or change from baseline). These choices were typically made to match the domain and scale used in the model and the model type (Markov or microsimulation). In addition, implementations differed in the level of premodel analysis conducted using the hypothetical trial data, such as estimating the ratio in mean change from baseline between the intervention and control arms, fitting a survival model to the individual patient-level data to obtain a hazard ratio, applying the relative effect on a continuous outcome as a relative risk (RR) reduction on a transition probability, or calibration. None of the included models used the hypothetical trial biomarker outcomes to implement a treatment effect. Discontinuation and waning varied in terms of being implemented at all (6 of 9 models) and in their timing and duration. See Table 2 for details.

Between-Model Variability

Disease progression in the control strategy and the difference with the intervention strategy varied across models. Mean time spent in MCl over a 10-year time horizon in the control strategy ranged from 2.6 to 5.2 years (see Fig. 1). The difference in MCl between intervention and control strategy ranged approximately 0.1 to 1.0 years (-0.8 to 0.3 for mild, -0.5 to 0.0 for moderate, -0.6 to 0.0 for severe; Fig. 2). The proportion in MCl over time in the control strategy varied between models, eg, range 0.16 to 0.60 at year 3 and 0.04 to 0.42 at year 5 (see Appendix Fig. 6.1 in Supplemental Materials found at https://doi.org/10.1 016/j.jval.2024.09.006). Two models (BASQDEM and RTI-HS) showed a relatively slower progression than other models in year 1 and faster progression for year 2 onward.

QALYs and costs in the control strategy ranged from 4.2 to 6.3 and from \$68 558 to \$451 421, respectively. The difference in QALYs and costs (excluding treatment costs) between intervention and control ranged from 0 to 0.6 and -\$66 897 to \$11 896, respectively (see Fig. 3). One model (RTI-HS) showed a relatively large cost saving and only one model (BASQDEM) showed small increased costs.

Sex differences (women minus men) in incremental 10-year time in MCI, cumulative QALYs, and cumulative costs ranged from 0.01 to 0.07, -0.01 to 0.05, and -\$13 147 to \$349 respectively, with the largest differences by 2 models (RTI-HS and BASQDEM).

Possible Causes for Variation

We performed additional post hoc analyses (see Appendix 6 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2 024.09.006) to assess the correlation among time gained in disease state (Fig. 2), QALY gains and differences in costs (Fig. 3), and implementation details (Table 2).

The variation in predicted gained time in MCI seemed relatively strongly related to the treatment effect estimate (see Appendix Fig. 6.2 in Supplemental Materials found at https://doi. org/10.1016/j.jval.2024.09.006). The trial outcomes chosen by the modeling groups from the benchmark hypothetical trial varied meaningfully. This resulted in the use of different treatment effect estimates (ie, the relative effect of the hypothetical benchmark intervention expressed as the ratio in average change from baseline between control and intervention arm, being CDR-SB = 0.70, CDR-global = 0.93, MMSE = 0.65, and FAQ = 0.73, and expressed as the hazard ratio from trial individual patient-level data Cox survival analysis being CDR-SB = 0.80). Nevertheless, 5 models that selected the same treatment effect outcome (CDR-SB) deviated relatively widely in their gained time in MCI (FEM = 0.3 years, IPECAD = 0.4, Biogen = 0.6, Bedrejo = 0.6, and BASQDEM = 0.9). Among other factors described below, the difference between 2 of these models (IPECAD and Biogen) could be explained by one of them (Biogen) implementing the trial ratio in change from baseline of 0.70 directly as an RR for transitioning to dementia, whereas the other (IPECAD) fitted a survival model that produced a hazard ratio of 0.80 implemented as an RR for transitioning to dementia. The difference between 2 other models (IPECAD and Bedrejo), which both applied the same RR of 0.80, could be explained by one of them (Bedrejo) applying a relatively small treatment effect waning. The similarity between 2 other models (Biogen and Bedrejo) could be explained by different assumptions canceling one another out: one of them (Biogen) applied a higher treatment effect (RR 0.70) but with waning whereas the other (Bedrejo) applied a lower treatment effect (RR 0.80) but with limited waning. The difference between these models and

Table 2. Implementation of treatment effect.

| Group | Markov type | | | | | Microsimulation type | | | |
|---|---|---|--|---|---|---|---|--|--|
| Participant | Biogen | CPEC | IPECAD | Bedrejo and colleagues | SveDem | BASQDEM | FEM | RTI-HS | MISCAN- Dementia |
| Implementation | | | | | | | | | |
| Choice of outcome(s) | CDR-SB, change from baseline | MMSE, change from baseline | CDR-SB, time to dementia | CDR-SB, time to dementia | CDR-global, change from baseline | CDR-SB, change from baseline | CDR-SB, change from baseline | MMSE and FAQ, changes from baseline | CDR-global, change from baseline |
| Rationale | Health states in the model defined using CDR-SB ranges | Health states in the model defined using MMSE ranges | Measure captures multiple domains and is sensitive in early AD. | Health states in the model defined using CDR-SB ranges | Transition to dementia is clinically relevant; aligns with use of synthetic control arm for MCI natural history | Time to dementia in the model relies on equation for CDR-SB progression over time. | Cognitive states in the model are defined using CDR-SB, which also contributes to the staging of dementia. | Aligned with the cognitive (MMSE) and functional (DAD) domains in the model (with mapping from FAQ to DAD) | Appropriate for the model's time to dementia approach; aligns with use of synthetic control arm for MCI natural history |
| Estimate(s) used/ metric used | Relative difference in CDR-SB change from baseline at 18 months | Relative difference in MMSE change from baseline at 18 months | HR for time to CDR-SB ≥4.5 events | [Same as IPECAD] | Relative difference in proportions transitioning to CDR-global 1 at 18 months | CDR-SB change from baseline at each time point in both arms | Relative difference in CDR-SB change from baseline at 18 months | Relative reduction in MMSE and FAQ change from baseline at 18 months | CDR-global change from baseline at 18 months |
| Relative effect observed in trial of the chosen outcome | HR = 0.70 | RR = 0.65 | HR = 0.80 | HR = 0.80 | RR = 0.93 | The regression mixed model contains a beta to include the trial arm | HR = 0.70 | Relative decline = 0.65 (MMSE) and 0.73 (FAQ) | RR = 0.93 |
| Synthetic trial individual-level data analysis, if applicable | None | None | Cox survival model using time to CDR-SB ≥4.5 events | [Same as IPECAD] | None* | Mixed regression model fitted to CDR-SB individual- level synthetic trial data | None | None | None |
| Implementation in model | Applied as a HR to transition probabilities from MCI to mild dementia and from mild dementia to moderate dementia | Applied as an RR to transition probabilities from MCI to mild dementia and from mild dementia to moderate dementia | HR applied directly to transition probability from MCI to mild dementia; calibration required to model treatment effect in mild dementia | Applied as a HR to transition probabilities from MCI to mild dementia and from mild dementia to moderate dementia | Applied as an RR to transition probabilities from MCI to mild dementia and from mild dementia to moderate dementia | Parameter in mixed regression model to reflect difference between control and intervention arm CDR-SB | Applied as an HR to transition probabilities from MCI to mild dementia and from mild dementia to moderate dementia | Relative reduction in MMSE and FAQ change from baseline applied to MMSE and DAD annual rates of decline, assuming a linear mapping between FAQ and DAD | Multiplicative factor for time in MCI calibrated to the CDR- global transitions at 18 months for intervention arm; calibrated factor also applied to time in mild dementia |
| Extrapolation beyond 18-month synthetic trial data | | | | | | | | | |
| Extrapolation under treatment [†] | Effectiveness maintained in MCI and mild dementia | Effectiveness maintained in MCI and mild dementia | Effectiveness maintained in MCI and mild dementia | Effectiveness maintained in MCI and mild dementia | Effectiveness maintained in MCI and mild dementia | Effectiveness maintained in MCI | Effectiveness maintained in MCl and mild dementia | Effectiveness maintained in MCI and mild dementia | Effectiveness maintained in MCI and mild dementia |
| Type of extrapolation after treatment discontinuation [†] | Disease- modifying: revert to natural history progression (gains maintained) | Disease- modifying: revert to natural history progression (gains maintained) | Disease- modifying: revert to natural history progression (gains maintained) | Disease- modifying: revert to natural history progression (gains maintained) | Disease- modifying: revert to natural history progression (gains maintained) | Partly disease- modifying: revert to natural history progression (gains maintained); partly symptomatic: revert to natural history state (gains lost) | Disease- modifying: revert to natural history progression (gains maintained) | Disease- modifying: revert to natural history progression (gains maintained) | Disease- modifying: revert to natural history progression; gains maintained |

Table 2. Continued

| Group | Markov type | | | | | | Microsimulation type | | | |
|---|---|--|--|---|---|---|-----------------------------------|--|---|--|
| Participant | Biogen | CPEC | IPECAD | Bedrejo and colleagues | SveDem | BASQDEM | FEM | RTI-HS | MISCAN- Dementia | |
| Waning description | Waning of 5% of effect lost per year starting in year 2 | No waning implemented | Waning of 5% of effect lost per year in MCI starting in year 2; time-dependent waning not applied to calibrated multi- domain effect in mild dementia, opting instead for a constant waning factor set to the 10-year average of the MCI waning | Waning was assumed to add up to 15% in the transition from mild to moderate, implemented as 0.80-(1-0.15) | Waning of 5% of effect lost per year starting in year 2 | No waning implemented | No waning implemented | Waning of 5% of effect lost per year starting in year 2 | 5% of patients lose effect per year; implemented by proportionally scaling the incremental time in MCI and mild dementia dependent on the timing of waning | |
| Discontinuation | | | | | | | | | | |
| Description | 10% over 18 months from synthetic trial intervention group; converted to 6.8% per year in all years | Not implemented | 10% in year 1; further discontinuation at 4% per year in all subsequent years | 10% in year 1 only | 10% in year 1 only | Not implemented | Not implemented | 10% over 18 months from synthetic trial intervention group; converted to 6.8% per year in all years | 10% over 18 months from synthetic trial intervention group; no discontinuation assumed after 18 months | |
| Stopping rule | At moderate dementia | At moderate dementia | At moderate dementia | At moderate dementia | At moderate dementia | After a maximum of 10 years on treatment | At moderate dementia | At moderate dementia | At moderate dementia | |
| Adverse events | | | | | | | | | | |
| Were adverse events implemented and, if so, how? | Implemented ARIA events in year 1 only and fall and headache event rates per year from synthetic trial intervention and control data; costs and disutilities assigned per event | AEs and ARIA not explicitly considered | Accounted for ARIA events in year 1 only from incremental synthetic trial event data; implemented via a single disutility and cost | Disutility of 0.14 in year 1 was applied for ARIA- E to the proportion of the baseline population, estimate was obtained from the symptomatic ARIA-E difference between synthetic data intervention and control; assumed 3 MRI scans for each individual in year 1 | AEs accounted for via a cost of \$1000 in year 1 only | AEs and ARIA not considered | AEs and ARIA not considered | AEs and ARIA not explicitly considered; assumed to be reflected in treatment discontinuation rate | AEs and ARIA not considered | |
| Model outcomes | | | | | | | | | | |
| Proportion in dementia year 1 (conditional on survival) in control strategy | 15% | 27% | 22% | 23% | 14% | 3% | 14% | 10% | 14% | |
| Ratio of proportion in dementia between control | 0.72 | 0.65 | 0.83 | 0.80 | 0.94 | 0.43 | 0.71 | 0.18 | 0.89 | |

and intervention

AD indicates Alzheimer's disease; AE, adverse event; ARIA, amyloid-related imaging abnormality; CDR-SB, clinical dementia rating – sum of boxes; DAD, disability assessment for dementia; FAQ, functional activities questionnaire; HR, hazard ratio; MCI, mild cognitive impairment; MMSE, mini-mental state examination; MRI, magnetic resonance imaging; RR, relative risk.

*An attempt was made, but the result was considered not valid and therefore not used.

[†]Extrapolation categories:

Absolute control counterfactual / gains lost / symptomatic: revert to control strategy counterfactual (ie, natural progression) in terms of absolute symptoms at same time.

Relative counterfactual / gains maintained / disease-modifying: revert to control strategy counterfactual (ie, natural progression) in terms of rate of decline at same state.

• Effectiveness maintained / cure: effect of the treatment remains applicable.



Figure 1. Model prediction of time spent in states over a 10-year time horizon for the control strategy, for men (left) and women (right).

another model (BASQDEM) is possibly due to faster natural progression creating a larger room for improvement as described below (and vice versa for FEM due to slower progression creating a smaller room for improvement).

Predicted gained time in MCI seemed related to natural disease history (ie, time spent in MCI in the control strategy) (see Appendix Fig. 6.3 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2024.09.006), although somewhat less than the relation with "choice of outcome(s)" as described earlier. Spending less time in MCI implies a higher risk of dementia onset, which generates a greater room for improvement (ie, a larger proportion to be prevented). Nevertheless, 4 models (BASQDEM, IPECAD, Bedrejo, and MISCAN) showed a relatively large variation, which could be explained by the treatment effect (increasing respectively with their deviation); the same accounts for the difference between 2 models (Sve-Dem and Biogen). One model (BASQDEM) showed almost no time in moderate or severe states leaving less room to reduce

time spent in these states, which are associated with low quality of life and high cost.

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QALY gains could be explained by the gained time in MCI given that they seemed relatively strongly related (see Appendix Fig. 6.4 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2 024.09.006). Nevertheless, 4 models (BASQDEM, CPEC, Biogen, and Bedrejo) showed a relatively large range in QALY gain at similar gained time in MCI. The difference between 2 models (Biogen and Bedrejo) could be explained by one of them (Bedrejo) applying lower waning in mild dementia and therefore gaining more QALYs from the treatment in that disease stage. The differences between the 2 other models (BASQDEM and CPEC) were difficult to explain and possibly related to differences in dementia natural history between them.

Differences in costs (excluding treatment costs) could partly be explained by the gained time in MCI given that they seemed somewhat related (see Appendix Fig. 6.5 in Supplemental Materials found at https://doi.org/10.1016/j.jval.2024.09.006).



Figure 2. Model prediction of difference (in terms of intervention minus control strategy) in time spent in states over a 10-year time horizon for men (left) and women (right).



Figure 3. Model prediction of incremental QALYs and incremental costs (without treatment costs) for men (left) and woman (right) over a restricted 10-year time horizon.

Nevertheless, 3 models with similar gained time in MCI (BASQ-DEM, CPEC, RTI-HS) showed a relatively large variation in costs. This could be explained by one of them (BASQDEM) "trading" time in mild dementia for time in MCI and another (RTI-HS) "trading"

Table 3. Suggested standard template to report predictions of disease progression from future AD simulation models: proportion of persons in each state over time in the control and intervention strategy over a 25-year time horizon separately for men and women; undiscounted and no half-cycle correction (ie, proportion at end of year; eg, at year 2, the proportion at day 365×2 is to be provided).

| Year | MCI | Mild dementia | Moderate dementia | Severe dementia | Death |
|--------------|-----|------------------|----------------------|--------------------|-------|
| Control | | | | | |
| 0 | 1 | 0 | 0 | 0 | 0 |
| 1 | | | | | |
| 2 | | | | | |
| 3 | | | | | |
| | | | | | |
| 25 | | | | | |
| Intervention | | | | | |
| 0 | 1 | 0 | 0 | 0 | 0 |
| 1 | | | | | |
| 2 | | | | | |
| 3 | | | | | |
| | | | | | |
| 25 | | | | | |
| | | | | | |

AD indicates Alzheimer's disease; MCI, mild cognitive impairment.

time in moderate and severe for time in MCI and mild. With higher costs in more severe stages, the benefits of "trading" those stages are higher. In one model (BASQDEM), the costs related to gained life-years spent in MCI were likely not offset by savings related to prevention of spending life-years in mild and moderate dementia. This balanced effect in costs does not apply to QALYs where all additional life-years result in QALY increase rather than additional care costs in life-years gained. Other causes could relate to cost input estimates used or impact on mortality making it difficult to pinpoint the precise cause for variation in costs.

We explored various other factors to explain model differences. Microsimulation-type versus Markov-type differences were not uniformly related to gained time in MCI (see Fig. 2) and varied relatively largely in QALY gains and costs (excluding treatment costs) (see Fig. 3). The differences seem to be better explained by the other factors described earlier. Sex differences showed slightly higher benefits in women in gained time in MCI (see Fig. 2) and health-economic predictions (see Fig. 3), most likely related to a higher life expectancy providing a larger room for treatment benefit.

None of the models used the benchmark biomarker results, none selected the composite outcome (AD composite score), none implemented a direct effect on mortality (ie, unrelated to disease state), and all assumed the treatment to be (partly) of disease-modifying nature in terms of reverting to natural history progression rate with treatment gains maintained once treatment is ceased. These factors remain undetermined as possible factors that could cause variation across model predictions.

Discussion

This study cross-compared the predictions from 9 AD decisionanalytic models that started in MCI and independently implemented efficacy estimates from the same benchmark hypothetical trial for an AD treatment. Variation was found in model Table 4. Suggested standard template to report health-economic predictions from future AD simulation models: cumulative outcomes over 10-year time horizon separately for men and women (mean per person).

| Outcome | Control | Intervention |
|---|---------|--------------|
| Time (years) on intervention (undiscounted) | | |
| Time in full-time/institutionalized care (undiscounted) | | |
| Direct costs medical (discounted and undiscounted) | | |
| Direct costs nonmedical (discounted and undiscounted) | | |
| Patient QALY (discounted and undiscounted) | | |
| Informal carer QALY (discounted and undiscounted) | | |
| AD indicates Alzheimer's disease; QALY, quality-adjusted life-year. | | |

predictions of natural disease history, time gained in MCI, QALYs, and costs.

Our results suggest the following cascade of factors that explained differences in model predictions: model design drove the choice of the outcome selected from trial evidence (eg, CDR-SB, MMSE), which was associated with the relative effect size from the trial evidence (7%-35%), which together with natural history (faster progression creating a larger room for improvement) and assumptions on waning and discontinuation (eg, effectiveness maintained under and after treatment, reversion to natural history progression after discontinuation) drove the gained time in less severe disease states (particularly MCI), which drove the QALY gain (by both more time in less severe states and increased life expectancy) and costs (although less straightforward as cost-savings due to more time in less severe states are offset by additional costs in life-years gained). Time spent on treatment and model type (Markov or microsimulation) did not seem to explain model differences. We note treatment costs were not included, although these are expected to be influenced by time on treatment.

Although models varied in their choice of treatment effect outcome, the modeling groups' rationales for linking their selected outcomes with the approaches used to model natural disease progression suggested face validity. However, we think applying the ratio between mean change from baseline on continuous outcomes such as CDR-SB and MMSE as an RR of conversion from MCI to dementia or from mild to moderate dementia (as applied by Biogen and CPEC models in this analysis) is uncertain due to limited empirical evidence. However, this choice might have been driven by limited resources not allowing exploration of alternative options that are likely more time consuming such as analyzing the synthetic individuallevel trial data or addressing this choice in sensitivity analysis.

The observed variation in natural history in the current IPECAD workshop cross-comparison study showed similarities to the variation in the previous IPECAD workshop cross-comparison in 2020.⁶ The previous workshop specified the benchmark as a 30% reduction in progression from MCI to dementia rather than a set of plausible trial outcomes with different relative effects in our current study. The range of time (years) spent in MCI (natural history) in our current study (2.4 to 5.2) was lower and wider than the range reported by the previous cross-comparison (3.4 to 5.6). We think this is due to an emphasis on the more specific amyloidconfirmed MCI starting population, which reflects a higher risk of dementia onset (ie, shorter time in MCI). The difference in time spent in MCI between intervention and control in our current study (0.1-1.0) was lower than the previous cross-comparison (approximately 0.3-1.7 with 1.1 being the second highest). We consider this is due to the relative treatment effect ranging from 7% to 35% compared with the uniform application of a 30% reduction in the previous cross-comparison.

A previous study focusing on structural uncertainty and crossvalidation of decision models in AD⁵¹ applied a hypothetical 30% reduction in disease progression from MCI to dementia using 3 different model types (1 Markov and 2 microsimulation) in combination with 4 data sources for natural progression, holding other inputs and assumptions constant. Incremental QALYs ranged from approximately 0.35 to 0.50 but when applying a 15% treatment effect instead of 30% it ranged from 0.15 to 0.25. We judge that these results compare favorably with the range in our current cross-comparison (0.05 to 0.55), anticipating this corresponds to a 7% to 35% treatment effect.

Recommendations for Reimbursement Agencies and Modelers

Based on our findings, we have 5 key recommendations for researchers and experts. First, for the evaluation of treatments for AD, we recommend that modelers address and discuss choices and assumptions in sensitivity analyses, especially those listed in Tables 1^{14-48} and 2 along with other assumptions highlighted by earlier cross-comparison and reviews, such as mortality⁶ and alternative data sources.^{52,53} We recommend that reimbursement agencies request the results of such analyses or test them themselves, for example, using one of the available AD open-source models.^{16,54,55} We note that our study did not produce the currently lacking empirical evidence on what choices and assumptions are correct, but merely the uncertainty around their choice.

Second, we recommend for modelers to report, and for reimbursement agencies to request, a standardized table reporting the proportion in different states at each year (timepoint), as designed in Tables 3 and 4. This adds to generic⁵⁶ and AD-specific⁵⁷ reporting recommendations on (dis)aggregated health-economic outcomes. For this table, we acknowledge the possible requirement of categorization and/or mapping⁵⁸ into disease stages being a potential source for variation in itself, but we consider it the optimal balance between the level of detail and outcomes most AD models can provide.

Third, we recommend that reimbursement agencies support the generation of long-term real-world evidence registries (incorporating both clinical and health-economic outcomes). Such evidence can be used for extrapolating treatment effects beyond short-term trial follow-up^{59,60} to reduce uncertainty related to assumptions on extrapolation in health-economic models.

Fourth, we invite modelers to submit their results to our IPECAD registry for open continuous cross-comparison (https://osf.io/jv85a/) in the format of Tables 3 and 4 as well as earlier Tables 1¹⁴⁻⁴⁸ and 2 to further support the transparency and credibility of AD models. We recommend this both for our benchmark scenario and for recent AD drug trials such as aducanumab,⁸ lecanemab,⁹ and donanemab¹⁰ to support ongoing model cross-validation in AD.

Fifth, because gained time in MCI was an important driver of health-economic model predictions, we recommend modelers to detail the MCI stage of a model in at least as much detail as the dementia stage, noting that earlier reviews recommended reflecting dementia as continuous disease progression on multiple domains.^{61,62} This could be done, for example, by explicitly modeling or using subpopulation-specific input estimates for cognitive status (eg, early/late or [non]amnestic cognitive impairment) and biomarker status ([ab]normal amyloid, [ab] normal neuronal injury), as provided by Vos et al.²³ Because "trading" time between less and more severe stages (as described earlier) seemed to drive health-economic model predictions, it remains important to reflect dementia stages in some detail.

Strengths/Limitations

Of the modeling groups invited to participate, 12 groups did not participate due to nonresponse and 10 groups could not participate due to insufficient resources, which limits the generalizability of our results. However, among the 12 nonresponders, most models were relatively old or reflected categorized disease progression states, a well-represented model type in our cross-comparison. Nevertheless, of the 8 models with limited resources to participate, one explicitly modeled biomarkers²⁴ and one was used for the evaluation of aducanumab and lecanemab,⁶³ which would likely have enriched the data of our cross-comparison.

The analyses described here omit treatment costs and hence do not inform cost-effectiveness estimates for these therapies. In addition, our benchmark scenario did not cover any potential diagnostic procedures, limiting generalizability to the assessment of real-world pharmacological interventions. We note the method to implement mortality did not differ between models but has previously been shown to affect health-economic model predictions.^{6,21,64,65}

Although we sought to maintain independence among the participating modeling groups, this was limited for SveDem given that it received technical model coding support and absent for Bedrejo given that it received recommendations on implementing the benchmark scenario from an IPECAD steering member (R.H., also part of the IPECAD modeling group). Some modeling groups participated in the original development of multiple models (R.H., W.L.H.) without being involved in the implementation of the benchmark scenario. In addition, some results were adjusted (RTI-HS) or submitted (FEM) after participation in the workshop.

Conclusions

Our benchmark scenario for a hypothetical set of trial results for an AD treatment was implemented independently in different ways across 9 distinct health-economic models and resulted in a relatively large variation in model predictions. Our results suggest the following cascade of effect: model design drove the choice of outcome, which was associated with relative effect size, which together with natural history (faster progression creating a larger room for improvement) and assumptions on waning and discontinuation drove the gained time in less severe disease states, which drove the QALY gain and costs. We recommend (1) addressing outcome choice and treatment effectiveness assumptions as part of sensitivity analysis, (2) using a standardized table for reporting AD model predictions, and (3) exploring the use of registries for future AD treatments measuring long-term disease progression to reduce the uncertainty of extrapolating trial results in health-economic models.

Author Disclosures

Author disclosure forms can be accessed below in the Supplemental Material section.

Supplemental Material

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Author Affiliations: Alzheimer Centre Limburg, Faculty of Health Medicine and Life Sciences, School for Mental Health and Neuroscience, Department of Psychiatry and Neuropsychology, Maastricht University, Maastricht, The Netherlands (Handels); Division of Neurogeriatrics, Department of Neurobiology, Care Sciences and Society, Karolinska Institutet, Solna, Sweden (Handels, Herring, Aye, Tate, Green, Gustavsson, Wimo, Winblad, Sköldunger, Jönsson); Health Economics, RTI Health Solutions, Research Triangle Park, NC, USA (Herring, Kamgar); Biogen Idec Ltd, Maidenhead, England, UK (Green); Quantify Research, Stockholm, Sweden (Gustavsson); Theme Inflammation and Aging, Karolinska University Hospital, Huddinge, Sweden (Winblad); Clinical Memory Research Unit, Department of Clinical Sciences, Lund University, Lund, Sweden (Raket); Community Health Sciences & O'Brien Institute of Public Health, University of Calgary, Calgary, Alberta, Canada (Stellick, Spackman); Health Economics, Policy and Innovation Institute, Masaryk University, Brno, Czech Republic (Hlávka); USC Price School of Public Policy and Schaeffer Center for Health Policy and Economics, Los Angeles, CA, USA (Hlávka, Wei); Basque Health Service (Osakidetza), Debagoiena Integrated Healthcare Organisation, Research Unit, Arrasate-Mondragón, Spain (Mar); Biogipuzkoa Health Research Institute, Donostia-San Sebastián, Spain (Mar); Biosistemak Institute for Health Service Research, Barakaldo, Spain (Mar); Faculty of Engineering, Electronics and Computing Department, Mondragon Unibertsitatea, Mondragon, Gipuzkoa, Spain (Soto-Gordoa); Department of Public Health, Erasmus MC, University Medical Center Rotterdam, Rotterdam, The Netherlands (de Kok, Brück); Care Policy and Evaluation Centre, London School of Economics, London, England, UK (Anderson); Biogen International GmbH, Baar, Switzerland (Pemberton-Ross, Urbich).

Correspondence: Ron Handels, PhD, Department of Psychiatry and Neuropsychology, Maastricht University, Universiteitssingel 40, 6200 MD, Maastricht, The Netherlands. Email: ron.handels@maastrichtuniversity.nl

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